

Gone with the Big Data:
Institutional Lender Demand for Private Information *

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Gone with the Big Data:

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ABSTRACT

I explore whether big-data sources can crowd out the value of private information acquired through lending relationships. Institutional lenders have been shown to exploit their access to borrowers' private information by trading on it in financial markets. As a shock to this advantage, I use the release of the satellite data of car counts in store parking lots of U.S. retailers. This data provides accurate and near-real-time signals of firm performance, which undermines the value of borrowers' private information obtained through syndicate participation. I find that once the satellite data becomes commercially available, institutional lenders are less likely to participate in syndicated loans. The effect is more pronounced when borrowers are opaque or disseminate private information to their lenders earlier and when the data predicts borrower performance more accurately. I also show that institutional lenders' reduced demand for private information leads to less favorable loan terms for borrowers.

Keywords: Debt Contract, Relationship Lending, Information Asymmetries, Institutional Investors, Informed Trading, Big Data, Satellite Images, Alternative Data, Fintech

JEL Codes: D82, G14, G21, G23, G30, M41

1. Introduction

Over the past two decades, the influx of nonbank institutional lenders as syndicated loan participants has driven the growth of the large corporate loan market (Ivashina and Sun 2011a; Jiang et al. 2010; Lim et al. 2014; Peyravan 2020).¹ The outstanding amount of syndicated institutional loans increased from \$100 billion in 2000 to \$1 trillion in 2018 (FDIC 2019). The migration of corporate credit risk to institutional lenders has been facilitated in part by low-interest environments and tighter banking regulations after the global financial crisis (Irani et al. 2021). Importantly, institutional lenders are not subject to stringent banking regulations and have been shown to exploit their private information advantages in the equity, bond, and credit derivatives markets by trading on borrowers' private information gained through lending relationships (e.g., Acharya and Johnson 2007; Bushman et al. 2010; Han and Zhou 2014; Haselmann et al. 2022).² Moreover, the informed trading opportunities embedded in the lending relationship generate economically significant profits (e.g., Ivashina and Sun 2011b; Massoud et al. 2011; Peyravan 2020). To shed more light on institutional lending, I examine whether institutional lenders' demand for valuable private information is a significant determinant of their participation in syndicated loans.

The primary empirical challenge in estimating the demand of institutional lenders for borrowers' private information is that lenders' information acquisition is not observable. To overcome this challenge, I take advantage of the availability of alternative data that undermines the value of borrowers' private information. Specifically, I use satellite image data from Orbital Insight that tracks the number of cars in store parking lots for a subset of U.S. retailers. This data

¹ Institutional lenders typically include investment banks, insurance and finance companies, mutual funds, pension funds, collateralized loan obligations (CLOs), private equity funds, and hedge funds.

² Over the course of a loan, borrowers regularly provide their lenders with non-public information, which includes monthly financial statements, covenant compliance certificates, amendment requests, and financial projections (Carrizosa and Ryan 2017; Standard & Poor's 2020).

has two important advantages. First, it provides valuable information about underlying firm performance (Kang et al. 2021a; Katona et al. 2021). Second, it is updated on a daily basis, allowing investors who purchase it to obtain timely updates on firm performance even before firms publicly disclose their performance. These two unique aspects of the satellite data are important in addressing my research question as they feature key aspects of the private information (i.e., early access to information about borrowers' performance) exploited by institutional lenders for insider trading.

When institutional lenders can access near-real-time information on a borrower's performance through this alternative source, the value of early access to borrowers' performance information through syndicate participation diminishes. Moreover, even if institutional lenders do not directly use the alternative data, the information advantages of these institutional lenders relative to other investors should also decline when other investors can take advantage of the satellite data, thus reducing incumbent institutional lenders' expected profits from their informed trading (Kyle 1985; Holden and Subrahmanyam 1992; Foster and Viswanathan 1996; Back et al. 2000; Akins et al. 2012; Katona et al. 2021). As a result, institutional lenders should have a lower demand for private information acquired through lending relationships, decreasing their incentives to extend loans to borrowers covered by the satellite data. Therefore, I predict that the probability of institutional lenders participating in a loan syndicate is lower when the satellite data on a borrower becomes commercially available.

To isolate the effect of changes in the value of borrowers' private information, I employ a difference-in-differences approach that compares the probability of institutional lenders participating in loans to firms with the satellite data coverage ("treatment borrowers") and to firms without such coverage ("control borrowers") before and after the initiation of the satellite data coverage. I focus on institutional lenders that engage in investment businesses, including

investment banking, asset management, private equity, and hedge-fund management. These businesses provide a platform for institutional lenders to extract benefits from timely access to value-relevant information about their borrowers. Furthermore, I require these institutional lenders not to be subsidiaries of bank holding companies, as banks are subject to greater regulatory scrutiny and typically have stronger internal controls, which may diminish institutional investors' opportunities for insider trading on borrowers' private information.³

Consistent with my prediction, I find that institutional lenders are less likely to issue loans to borrowers covered by the satellite data, controlling for borrower and loan characteristics as well as firm, quarter-of-loan-origination, credit-rating, and loan-type fixed effects. Economically, the probability that institutional lenders issue loans to treatment borrowers relative to control borrowers decreases by 10% in the coverage period after the release of the satellite data.

To assess the validity of the parallel-trend assumption in difference-in-differences estimation, I demonstrate that the probabilities that institutional lenders issue loans to treatment and control borrowers are not statistically different in the pre-coverage period before the satellite data is commercially available. To further mitigate concerns that the results may be affected by other confounding factors, I control for differences in observable characteristics across treatment and control borrowers, using an entropy balancing approach. Using this matching technique, I find consistent evidence that the satellite data coverage decreases institutional lending.

Next, I perform falsification tests using other types of institutional lenders that are unlikely to exploit early access to borrowers' private information for insider trading purposes. First, institutional lenders will have limited demand for private information when they do not have a trading operation and thus cannot exploit their information advantages. Second,

³ I consider U.S. bank holding companies to be those identified by the Financial Stability Board (FSB) as global systemically important banks (G-SIB), which include JP Morgan, Bank of America, Wells Fargo, Citigroup, Goldman Sachs, Morgan Stanley, Bank of NY Mellon, and State Street.

institutional lenders affiliated with large bank holding companies face higher regulatory costs than independent institutional lenders and typically have controls in place to prevent the transfer of sensitive borrower information from loan officers to traders in other investment divisions (Carey et al. 1998; Peyravan 2020; Kang et al. 2022). These lenders therefore are less likely to trade on borrowers' private information. Consistent with the limited information demand of bank-affiliated institutional lenders and institutional lenders without investment businesses, I find no evidence that these lenders have a lower probability of issuing loans to borrowers covered by the satellite data.

To reinforce institutional lenders' information demand mechanism, I conduct a number of cross-sectional tests. I predict the negative effect of the satellite data coverage on institutional lenders' participation to be more pronounced for lenders that had a higher demand for borrowers' private information in the pre-coverage period (i.e., lenders that likely participated in syndicated loans primarily to obtain borrower's private information). Following the initiation of the satellite data coverage, these lenders perceive participation in the syndicated loans as less valuable because the alternative data is likely to substitute, at least partially, for borrowers' private information. To test this prediction, I first conjecture that institutional lenders should have a higher information demand when borrowers are opaque. Opaque borrowers provide imprecise public information that encourages private information acquisition and informed trading (Diamond 1985; Bushman 1991; Kim and Verrecchia 1991).⁴ Second, I expect institutional lenders to exhibit a higher information demand when borrowers disseminate private information to their lenders earlier, as timely information is more valuable to the lenders' trading activities.⁵

⁴ I measure a borrower's information opacity based on its analyst coverage, issuance of earnings forecasts, and press releases.

⁵ I measure early dissemination of borrower information based on whether a borrower issues loans with a higher number of performance (income-statement based) covenants or obtains loans from reputable lead arrangers (Bushman et al. 2010, Bushman and Wittenberg-Moerman 2012; Christensen and Nikolaev 2012).

Consistent with the prediction, I show that the satellite data coverage reduces institutional lending to a greater extent when borrowers are opaque or disseminate private information to their lenders earlier.

I also examine whether the satellite data coverage has a greater effect when the data is more accurate in predicting borrower performance. Increased precision of an alternative source of information can further crowd out the value of private information acquired through lending relationships. Therefore, when the satellite data provides more precise forecasts of borrower performance, I expect institutional lenders to have a lower demand for private information obtained through loan participation. I indeed find the effect of the satellite data coverage on institutional lending to be stronger for borrowers for which the satellite data is more accurate.⁶

I further investigate whether institutional lenders' reduced information demand affects borrowers' credit outcomes. I find that when institutional lenders stop funding loans to borrowers in the coverage period, these borrowers pay higher interest rates, receive smaller loan amounts, and obtain loans with shorter maturities. Such unfavorable loan terms are consistent with the lower information demand leading to a decrease in credit supply for borrowers covered by the satellite data. These results suggest that institutional lenders' information demand is an important factor that shapes loan contractual terms.

Next, I perform additional analyses to explore whether the enhanced transparency resulting from the availability of the satellite data can offset these unfavorable loan terms by reducing costs of capital from other sources. My results indicate that the cost of capital for opaque borrowers, who attract a higher information demand from institutional lenders, does not decrease in the coverage period, consistent with prior evidence that unequal access to the satellite data

⁶ I measure the accuracy of the satellite data using high correlation between car-count signals and firm performance or lower variability of car-count signals across the firm's stores.

may exacerbate information asymmetry between sophisticated investors and individual investors (Katona et al. 2021). I also do not find that these borrowers issue additional equities in the coverage period.

Lastly, I explore whether institutional lenders' information demand, particularly for opaque borrowers, affects loan spreads in the coverage period. Prior studies in compensation literature suggest that the optimal contract lowers explicit wages paid to managers when they can also be compensated through the expected insider trading profits (Baiman and Verrecchia 1996; Roulstone 2003). Drawing from this literature, institutional lenders' information demand can exert downward pressure on interest spreads for loans to opaque borrowers who provide greater opportunities for insider trading. On the other hand, prior studies indicate that institutional lenders have greater bargaining power because they often serve as lenders of last resort and, primarily extend credit to riskier borrowers, leading to higher premiums than for otherwise identical bank loans (Taylor and Sansone 2006; Nandy and Shao 2008; Lim et al. 2014). Consistent with the institutional lenders' strong bargaining power, I find that, despite the potential downward pressure on interest spreads due to information demand, institutional lenders do not reduce interest spreads for loans issued to opaque borrowers in the coverage period.

This paper makes several contributions. First, I contribute to the growing literature on nonbank institutional lending. Recent studies document that institutional lenders trade on borrower information obtained from their lending relationships (Ivashina and Sun 2011b; Massoud et al. 2011; Peyravan 2020).⁷ Moreover, institutional lenders accelerate the speed of

⁷ Peyravan (2020), who primarily focuses on the insider trading activities of dual holders (institutional investors that simultaneously hold a firm's loan and equity), also finds that these investors are more likely to invest in equities of borrowers with weak financial reporting quality. While these findings imply that institutional lenders tend to pursue opaque borrowers, my study directly examines whether institutional lenders' demand for private information is an important determinant of their syndicate lending participation, using the satellite data coverage as a shock to these lenders' information advantage.

stock-price discovery, especially for borrowers with weak public disclosure (Bushman et al. 2010), and stimulate greater borrower voluntary disclosure (Peyravan and Wittenberg-Moerman 2022). While these studies primarily explore the consequences of institutional lender participation on capital markets, I demonstrate that the demand for valuable private information is an important factor for institutional lenders' decisions to participate in syndicated loans. Relatedly, I contribute to the literature on the effect of institutional lender participation on loan pricing by providing evidence that institutional lenders' lower information demand adversely affects loan terms for borrowers (Jiang et al. 2010; Ivashina and Sun 2011a; Lim et al. 2014).

Second, I contribute to the emerging literature on the role of alternative big-data sources that are used by a subset of sophisticated investors in capital markets. Prior studies find that big data is a useful supplementary source of information that affects price informativeness, managerial actions, and information asymmetries among investors (Jame et al. 2016; Zhu 2019; Kang et al. 2021a; Katona et al. 2021). Relatedly, there is a growing literature on the effects of machine learning and information-sharing technologies on lending decisions (e.g., Sutherland 2018; Costello et al. 2020; Kang et al. 2021c; Bartlett et al. 2022; Liu 2022; Chen et al. 2023; Minnis et al. 2023). However, there is little work on how the availability of alternative big-data sources affects credit market dynamics. Using a big data source available to a subset of sophisticated investors, I demonstrate that the availability of the data source undermines information demand of institutional lenders, reducing their supply of credit.

Finally, I also contribute to the nascent literature on the importance of noncredit sources of income in private lending. Prior studies show that relationship lenders are more likely to obtain mandates for their borrowers' security underwritings and M&A deals (e.g., Drucker and Puri 2005; Yasuda 2005). These cross-sold products typically generate substantial income and can enhance the profitability of lending relationships. Therefore, banks take borrowers' cross-selling

potential into account when initiating lending relationships (Kang et al. 2021b). I complement these studies by showing that the potential trading benefits embedded in the lending relationship can significantly influence institutional lenders' incentives to retain these relationships.

The next section presents the hypothesis development. Section 3 describes data and sample selection. Section 4 reports main results, and Section 5 concludes.

2. Background and Hypothesis Development

2.1 Satellite Imagery Data

Satellite images are photos of Earth's surface collected by remote sensing satellites operated by government programs or commercial entities. Because satellite images have detailed and high-spatial resolution, they are large in size and mostly in an unstructured format, often referred to as "big data." Recent advances in machine-learning and cloud-computing techniques have made it feasible to parse out vast quantities of satellite images across the globe and extract useful information from them each day, enabling investors to "explore the world in real time". Investors can receive real-time updates on various economic activities measured based on satellite images of store parking lots, manufacturing centers, oil refineries, petrochemical plants, agricultural land, and mining operations, among others. These data help, for example, gauge a country's fuel supply, predict crop yields, estimate damages from natural disasters, and track flows and disruptions along supply chains.

In this paper, I use satellite image data – provided by Orbital Insight – which tracks the number of cars in parking lots for a subset of publicly listed U.S. firms. Orbital Insight was founded in 2013 and commercially released the car-count data in the third quarter of 2015. At the end of each day, Orbital Insight collects satellite images from its various providers including Landsat (a joint program of NASA and U.S. Geological Survey), DigitalGlobe, Airbus, and Planet Labs. Once these satellite images are gathered, Orbital Insight counts the number of cars in each

parking lot using a proprietary computer-vision and machine-learning algorithm which includes procedures to enhance the accuracy of the car-count data. For example, if multiple stores share the same parking lot, the algorithm identifies the area of the parking lot in front of each store's entrance and records the number of cars specific to the store. In this case, the algorithm provides information on a contamination level of each store's car-count data based on the probability of inaccurately counting the number of cars. Moreover, the car count-data is adjusted based on the time stamp of the satellite images to ensure comparability over time. Satellite images taken outside of operating hours and for stores with covered parking lots are excluded. Orbital Insight provides the car-count data to its clients the following morning.

2.2 Related Literature and Hypothesis Development

Over the past two decades, nonbank lenders have played an increasing role in supplying credit to corporations. In the syndicated loan market, nonbank lenders' share has grown from 40% in 2000 to 60% in 2014 (Peyravan 2020). Low-interest environments and tighter banking regulations after the global financial crisis facilitated the migration of corporate credit risk to nonbank lenders (Irani et al. 2021). Recent studies examine characteristics of nonbank loans and show that they have higher interest spreads, flexible covenants, and are more likely to be secured while their borrowers are smaller, less profitable, and have fewer financing alternatives (Lim et al. 2014; Chernenko et al. 2019; Loumioti 2022).

Nonbank lenders include a growing number of institutional lenders – such as investment managers, hedge funds, private equity funds, and investment banks – that also engage in investment businesses in financial markets. As syndicate participants, they have access to borrowers' performance information before it is publicly disclosed to market participants. Over the course of a loan, borrowers typically provide information to lenders on a monthly basis – including financial performance updates, covenant compliance reports, amendment requests,

and financial projections—and allow lenders to visit their sites (Carrizosa and Ryan 2017; Standard & Poor’s 2020; Gustafson et al. 2021). Regulators, banks, and borrowers have expressed considerable concern about institutional lenders’ access to private information (e.g., SEC 2012; Standard & Poor’s 2020). Some participants decide to waive their right to access borrowers’ private information to address potential concern over insider trading (Amiraslani et al. 2023).

Despite U.S. laws prohibiting trading on material private information,⁸ prior studies find that institutional lenders exploit their information advantages by engaging in insider trading in the equity, bond, and credit derivatives markets (Acharya and Johnson 2007; Bushman et al. 2010; Ivashina and Sun 2011b; Massoud et al. 2011; Han and Zhou 2014; Peyravan 2020). Insider trading generates economically significant profits; for example, institutional lenders can make abnormal profits of around \$5 million by short-selling borrowers’ stocks during the 20-day window around negative credit events (Massoud et al. 2011) or achieve 5% to 8% excess annual returns by trading borrowers’ stocks (Ivashina and Sun 2011b; Peyravan 2020).

I examine whether the value of borrowers’ private information is an important determinant for institutional lenders’ incentives to have lending relationships with borrowers. Empirical evidence on this topic is limited because lenders’ acquisition of private information is not directly observable. Recent studies use Freedom of Information Act (FOIA) requests to identify private information acquisition. Glaeser et al. (2023) show that information asymmetry between managers and outsiders promotes private information acquisition measured by FOIA requests submitted to the U.S. Securities and Exchange Commission. Down et al. (2022) find that lead arrangers file FOIA requests to the Food and Drug Administration to obtain private information about their borrowers. While FOIA requests reveal non-public information, they are

⁸ The Securities Exchange Act of 1934 and the Insider Trading Sanctions Act of 1984 are two federal laws that regulate insider trading.

fulfilled with considerable delay, therefore may not be useful for institutional lenders' instantaneous trading activities. To overcome such empirical challenges, I take advantage of the satellite image data from Orbital Insight that provides daily updates on the number of cars in store parking lots. This data has two important advantages. First, the car-count data provides valuable information by accurately predicting firm performance (Kang et al. 2021a; Katona et al. 2021). Second, the data is updated daily; therefore, investors can purchase the data to obtain timely updates on firm performances before firms publicly disclose their performance. These two unique aspects of the data are also the key characteristics of the private information exploited by institutional lenders for their informed trading in financial markets.

When institutional lenders can access the satellite data providing timely information on borrowers' performance, the value of private information acquired through lending relationships declines. Moreover, when other investors can also acquire the satellite data and trade on timely information about firm performance, institutional lenders expect greater competition in financial markets. The competition among informed investors reduces expected profits from their informed trading and discourages private information acquisition (e.g., Holden and Subrahmanyam 1992; Foster and Viswanathan 1993, 1996; Back et al. 2000; Akin et al. 2012). Institutional lenders should therefore have a lower demand to acquire private information by extending loans to borrowers covered by the satellite data. Building on these arguments, I predict that the probability that institutional lenders participate in syndicated loans is lower after the satellite data on a borrower becomes commercially available.

However, a number of factors may confound this prediction. First, factors other than the value of private information can dominate institutional lenders' incentives to have lending relationships with borrowers. For example, prior studies suggest that institutional lenders pursue syndicated loans because they offer high interest rates (Lim et al. 2014). Second, the satellite data

may be less informative than what institutional lenders can directly learn through syndicate relationships. Third, the satellite data may complement rather than substitute for the private information of institutional lenders. For example, the satellite data may help institutional lenders better understand private information related to borrower performance, thereby facilitating informed trading (Kim and Verrecchia 1994; McNichols and Trueman 1994). Fourth, the costs of acquiring and processing the satellite data can be prohibitive to investors (e.g., Blankespoor et al. 2020). Therefore, institutional lenders may continue to demand early access to borrower information through lending relationships. For these reasons, whether the availability of the satellite data attenuates institutional lending remains an open question.

3. Sample, Data, and Descriptive Statistics

3.1 Data Sources and Sample Selection

I obtain loan characteristics from DealScan and borrower characteristics from Compustat and CRSP. I collect analyst coverage data from I/B/E/S, press release data from RavenPack, and borrower credit ratings from Compustat and Mergent FISD. Satellite data coverage and store-level car-count data are from Orbital Insight. I select loans issued to U.S. borrowers in the same industries as borrowers covered by the satellite data, resulting in 6,907 loan packages from 2011 through 2019. I eliminate borrowers with missing Compustat identifiers, resulting in 2,684 loan packages. I match this sample to Compustat and further eliminate loans with insufficient borrower and loan characteristics. The final sample contains 98 treatment borrowers with the satellite data coverage and 546 control borrowers without the data coverage, corresponding to 2,129 loan packages syndicated by 677 lenders.

To identify institutional lenders, I first classify lenders as either commercial bank lenders or nonbank lenders. Following Lim et al. (2014), I identify a lender as a *commercial bank lender* if its lender type in DealScan is “US Bank,” “Foreign Bank,” “Thrift/S&L,” “African bank,” “Asian-

Pacific Bank,” “Eastern Europe/Russian Bank,” “Middle Eastern Bank,” or “Western European Bank.” I also classify a lender as a commercial bank lender if its SIC 4-digit code is between 6011 and 6082 or is 6712 or 6719. For each lender identified as commercial bank, I manually check whether the lender mainly engages in commercial banking business and exclude those that do not mainly accept deposits and extend individual or business loans.⁹ I classify all remaining lenders as *nonbank lenders*.

I further classify nonbank lenders based on whether or not they are affiliated with bank holding companies, using business descriptions from company websites, annual reports, Bloomberg, and Capital IQ. I classify a lender as a *bank-affiliated-institutional lender* if it is a subsidiary of a U.S. bank holding company, and otherwise as an *independent institutional lender*. For each *independent institutional lender* that is not affiliated with banks, I identify whether it engages in investment businesses, based on DealScan lender type. An independent institutional lender is considered to have an investment operation if its lender type in DealScan is “Inst. Invest. Prime Fd,” “Inst. Invest. Prime Hedge Fd,” “Inst. Invest. Prime CDO,” “Investment Bank,” “Mutual Fund,” or “Distressed (Vulture) Fund.” In addition, I check each lender’s business description to determine whether it engages in investment businesses, including investment banking, asset management, private equity, and hedge fund management.¹⁰ Finally, I classify the remaining lenders as independent institutional lenders that do not engage in investment businesses. These lenders include captive finance companies, lease companies, and farm credit institutions.

3.2 Descriptive Statistics

⁹ For example, I exclude from the commercial bank lenders category Goldman Sachs Group, ORIX USA Corp, and Pilgrim Group.

¹⁰ I focus on the independent institutional lenders that engage in investment businesses.

Table 1 presents descriptive statistics of the main variables used in the analyses. Of the loans in the sample, 16.5% are issued with at least one independent institutional lender engaging in investment businesses (*Inst. Lender*);¹¹ 33.2% are issued after the third quarter of 2015, when the satellite data becomes commercially available (*Post*). Treatment borrowers obtain 19.4% of sample loans (*Treatment Firm*). The sample borrowers are relatively large (*Assets*) and have an average leverage ratio (*Leverage*) of 0.348, a market-to-book ratio (*MTB*) of 3.23, a mean sales growth (*Sales Growth*) of 0.148, a mean interest coverage ratio (*Interest Coverage*) of 65.5.¹² They also have an average return on asset (*ROA*) of 0.027, an average Altman Z-score (*Altman Z*) of 3.546, an average age (*Age*) of 24 years, and an average stock return before the loan issuance (*Past Return*) of 5.5%. With respect to loan characteristics, the mean loan size (*Amounts*) is relatively large (USD \$406 million), the average maturity (*Maturity*) is approximately four years, and the average all-in-drawn spread (*Interest Spread*) is 195 bps. Around 50% of loans are secured (*Secured*) and 9% of sample loans have a guarantor (*Guarantor*). Detailed variable definitions are reported in Appendix A.

4. Research Design and Empirical Results

I organize my empirical analyses as follows. First, I examine the effect of the satellite data coverage on institutional lender participation in syndicated loans. Next, I explore the information demand channel by investigating whether the observed effect is stronger when institutional lenders are expected to have a higher demand for borrowers' private information in the pre-coverage period. Lastly, I examine whether institutional lenders' information demand affects borrowers' credit outcomes.

¹¹ Note that 24.8% and 5.9% of the loans in the sample are issued with, respectively, at least one bank-affiliated institutional lender (*Inst. Lender Bank Affiliated*) and at least one independent institutional lender without investment operations (*Inst. Lender No Investment*).

¹² The median value of *Interest Coverage* is 8.4. Main results are robust to winsorizing it at the 95% level (with its mean and SD of 22.1 and 39.09, respectively).

4.1 Satellite Image Data and Institutional Lending

I begin my analyses by investigating whether the probability that institutional lenders issue a loan is lower after the satellite data on a borrower becomes commercially available. My empirical strategy exploits the fact that a subset of U.S. borrowers is covered by the satellite data after the third quarter of 2015.¹³ I employ a difference-in-differences analysis using control borrowers in the same industries (SIC 4-digit) as the treatment borrowers.¹⁴ Specifically, I estimate the following model:

$$\text{Inst. Lender} = \beta_0 + \beta_1 \text{Treatment Firm} \times \text{Post} + \text{Controls} + \text{Fixed Effects} + \varepsilon, \quad (1)$$

In Model (1), the dependent variable (*Inst. Lender*) equals 1 if the loan is issued with at least one institutional lender that is not affiliated with a bank holding company and is engaged in investment businesses (hereafter, “institutional lender”), and 0 otherwise. The variable of interest is *Treatment Firm* \times *Post*, where *Treatment Firm* equals 1 if the borrower is covered by the satellite data after the data becomes commercially available (and 0 otherwise) and *Post* equals 1 if the loan is issued after the initiation of the satellite data coverage (and 0 otherwise). If institutional lenders are less likely to participate in loans to borrowers being tracked by the satellite data, I expect a negative and significant coefficient on *Treatment Firm* \times *Post*.

I control for borrower characteristics that can influence institutional lending decisions, which include a borrower’s size (*Assets*), liquidity (*Current Ratio*), leverage (*Leverage*), market price (*MTB*), sales growth (*Sales Growth*), interest coverage (*Interest Coverage*), profitability (*ROA*), credit risk (*Altman Z*), age (*Age*), and stock performance (*Past Return*). I also control for loan characteristics, including loan amount (*Amounts*) and, maturity (*Maturity*) and whether a loan is secured (*Secured*) or has a guarantor (*Guarantor*). I include firm fixed effects to control for

¹³ Note that firms have no control over whether they are covered by these data.

¹⁴ Main results are robust to using SIC 3-digit industries.

unobservable time-invariant characteristics of each firm and year fixed effects to control for time-varying factors common to all sample firms.¹⁵ I estimate Model (1) using a logit and an OLS model. In the OLS model, I substitute year fixed effects with year–quarter fixed effects and further include credit-rating and loan-type fixed effects.¹⁶ I cluster standard errors at the firm level.

I present my main findings in Table 2. Panel A reports the results of univariate tests. I find that when investors can access the satellite data ($Post = 1$), 8.7% of loans to treatment borrowers are issued with institutional lenders, compared to 19.4% for control borrowers. In contrast, when the satellite data is not available to investors ($Post = 0$), 16.0% of loans to treatment borrowers are issued with institutional lenders, compared to 16.3% for control borrowers. The difference-in-differences $((8.7\% - 19.4\%) - (16.0\% - 16.3\%))$ is statistically significant at the 1% level. This evidence is consistent with my prediction that institutional lenders are less likely to issue loans to borrowers covered by the satellite data.

Next, I report estimation results of Model (1) in Panel B of Table 2. In Column 1 (Columns 2 and 3), I employ a logit model (OLS models). I find a negative and significant coefficient on $Treatment\ Firm \times Post$ for all specifications.¹⁷ Economically, the probability of institutional lenders issuing a loan to a treatment borrower, relative to a control borrower, is 13.8% lower in the coverage period. I measure economic significance based on the OLS specification in Column 3, where I include firm, year–quarter, credit-rating, and loan-type fixed effects. These findings

¹⁵ I include year fixed effects, using indicator variables for the trailing 12 months ending in September of each year, to ensure that $Post$ does not have within-year variance. Thus, the coefficients on both $Treatment\ Firm$ and $Post$ are not estimated because they are perfectly collinear with year and firm fixed effects.

¹⁶ Due to issues regarding a large number of fixed effects in nonlinear models (e.g., Maddala 1987; Greene 2004), I include only year and firm fixed effects in the logit model.

¹⁷ With respect to controls, the negative and significant coefficients on ROA and $Past\ Return$ suggest that borrowers with higher profitability or higher prior buy-and-hold return are less likely to obtain loans from institutional lenders. The negative and significant coefficient on $Altman\ Z$ indicates that borrowers with higher credit risk attract institutional lenders. Institutional lenders are also more likely to participate in secured loans ($Secured$).

reinforce my prediction that the probability of institutional lenders issuing a loan is lower for borrowers being tracked by satellite data.¹⁸

The key identifying assumption of the difference-in-differences analysis is the parallel-trend assumption that institutional lending trends would be the same for both treatment and control borrowers in the absence of the satellite data coverage. In other words, it assumes that control borrowers provide the appropriate counterfactual of the trend that treatment borrowers would have followed if they had not been covered by the satellite data (Angrist and Pischke 2008). To examine whether the parallel-trend assumption holds, I estimate the following model:

$$\begin{aligned}
 \text{Inst. Lender} = & \beta_0 + \beta_1 \text{Treatment Firm} \times \text{Trend}_{t=-3,-4} + \beta_2 \text{Treatment Firm} \times \text{Trend}_{t=1,2} + \\
 & \beta_3 \text{Treatment Firm} \times \text{Trend}_{t=3,4} + \text{Controls} + \text{Fixed Effects} + \varepsilon,
 \end{aligned} \tag{2}$$

In Model (2), I replace *Treatment Firm* × *Post* in Model (1) with separate interaction variables between *Treatment Firm* and trend variables, each one of which equals 1 for every two-year sample period before and after the initiation of the satellite data coverage (and 0 otherwise). I exclude from the trend variables the last two-year period immediately before the release of the satellite data (from the fourth quarter of 2013 to the third quarter of 2015); therefore, this period serves as a benchmark period. Figure 1 graphically depicts the estimation results of Model (2). I find that the counterfactual treatment effect in the pre-coverage period (the coefficient on *Treatment Firm* × *Trend*_{t=-3,-4}) is statistically indistinguishable from that in the benchmark period,

¹⁸ I employ several alternative specifications to ensure that my results are not sensitive to research design choices. First, I re-estimate Model (1) using continuous dependent variables of the proportion (%) or the number of institutional lenders in the loan package and find robust results (reported in Panel A of Appendix B). Also, a subset of treatment borrowers in my sample are covered by the satellite data provided by RS Metrics. Main results are robust to using *Post RM* as the main variable of interest, which equals 1 if the loan is issued after the satellite data from either RS Metrics or Orbital Insight become commercially available, and 0 otherwise (results are tabulated in Panel B of Appendix B). Moreover, main results continue to hold when I exclude periods after SafeGraph, a major provider of mobile GPS location data, released its foot-traffic data in 2018 (results are tabulated in Panel C of Appendix B).

while treatment effects in the coverage period (the coefficients on $Treatment\ Firm \times Trend_{t=1,2}$ and $Treatment\ Firm \times Trend_{t=3,4}$) are significantly different from that in the benchmark period. These results provide support for the parallel-trend assumption.¹⁹

While firms cannot self-select to be covered by the satellite data, I recognize other factors that may confound my results. For example, if treatment and control borrowers differ in many dimensions, satellite data coverage may be endogenous with respect to these differences. In Panel E of Appendix B, I compare firm characteristics of treatment and control borrowers. Treatment borrowers are more profitable and older while they exhibit lower sales growth and have lower credit risk.²⁰ Although I control for time-invariant firm characteristics by including firm fixed effects in all analyses, I further alleviate this concern by employing the entropy balancing approach. This matching technique achieves covariate balance between treatment and control observations by re-weighting control observations ensuring that the mean and variance are identical across the matching variables for both treatment and control samples. Moreover, entropy balancing reduces bias from nonlinear relationships between observable characteristics and the dependent variable (Hainmueller 2012; McMullin and Schonberger 2020). In Panel C, Table 2, I present the estimation results using the entropy balanced sample and continue to find a negative and significant coefficient on $Treatment\ Firm \times Post$, consistent with the satellite data coverage curbing institutional lending.²¹

¹⁹ I also plot the probability of institutional lending for treatment and control borrowers separately during the sample period (untabulated). I visually check that these univariate trends do not indicate a violation of the pre-trend assumption. In addition, I restrict the main sample to the pre-coverage period and estimate Model (1) after interacting $Treatment\ Firm$ with a continuous trend variable of $Year$ (the year of loan issuance). As shown in Panel D of Appendix B, the coefficients on $Treatment\ Firm \times Year$ are not significant across analyses, which supports the parallel-trend assumption.

²⁰ I also examine whether characteristics of the treatment borrowers change in the post-coverage period relative to those of the control borrowers, which can potentially drive the main results. Using borrower characteristics such as assets size, credit rating, leverage, interest coverage, MTB, and earnings guidance, I find no such evidence (untabulated).

²¹ I check covariate balance of the entropy balanced sample in Panel E of Appendix B.

4.2 Falsification Test

In this section, I perform falsification tests to further support institutional lenders' information demand as a mechanism. I suggest that when the satellite data provides accurate and near-real-time signals on firm performance, institutional lenders have a lower demand to acquire borrowers' private information for insider trading purposes. Therefore, if the information demand is instrumental to the relationship between the satellite data coverage and institutional lending, my main results should not hold or at least be much weaker for other types of lender that are unlikely to exploit early access to borrowers' private information by engaging in informed trading.

Institutional lenders affiliated with bank holding companies are less likely to trade on borrower information obtained through lending relationships because these lenders are subject to stringent banking regulation and face higher regulatory oversight (Carey et al. 1998; Peyravan 2020). Moreover, bank-affiliated lenders tend to be larger organizations with controls in place to prevent the transfer of sensitive borrower information from loan officers to traders in other investment divisions who may exploit it (Carey et al. 1998; Peyravan 2020; Peyravan and Wittenberg-Moerman 2022). Therefore, I focus on loans issued with institutional lenders that are subsidiaries of bank holding companies. *Inst. Lender Bank Affiliated* equals 1 if the loan is issued with at least one bank-affiliated institutional lender but is not issued with an independent institutional lender (and 0 otherwise).

To exploit private information advantages for potential insider trading, lenders need to engage in investment businesses which can provide a platform to extract benefits using value-relevant information about their borrowers. Using loans issued with independent institutional lenders, I further identify those issued with lenders that do not engage in investment businesses.

Inst. Lender No Investment equals 1 if the loan is issued with at least one independent institutional lender that does *not* engage in investment businesses but is not issued with an independent institutional lender that does engage in investment businesses (and 0 otherwise).

I perform the falsification test by re-estimating Model (1) with each of these variables as the dependent variable. Panel A of Table 3 reports the estimation results. Consistent with my prediction, I failed to find a significant coefficient on *Treatment Firm* × *Post* across all specifications for which either *Inst. Lender Bank Affiliated* or *Inst. Lender No Investment* is the dependent variable. Next, I re-estimate Model (1) using a multinomial logit model. For this analysis, I create a dependent variable that takes the value of 1 if *Inst. Lender Bank Affiliated* equals 1, 2 if *Inst. Lender No Investment* equals 1, and, 3 if *Inst. Lender* equals 1 (and 0 otherwise). As reported in Panel B of Table 3, I failed to find significant coefficients on *Treatment Firm* × *Post* when the dependent variable equals 1 or 2, which suggests that the satellite data coverage does not affect loans issued with bank-affiliated institutional lenders or by independent institutional lenders without investment operations. These results are consistent with these lender types having low demand for borrowers' private information.

4.3 Institutional Lenders' Information Demand

To further support the information demand mechanism, I investigate whether the effect of the satellite data coverage on institutional lending is more pronounced if institutional lenders had a higher demand for borrowers' private information in the pre-coverage period. While higher information demand can stimulate institutional lenders to participate in syndicated loans, this participation becomes less valuable in the coverage period when the satellite data substitutes, at least partially, for borrowers' private information. Thus, I predict the effect of the satellite data coverage on institutional lending to be stronger for borrowers that attracted higher information

demand from institutional lenders before the initiation of the satellite data coverage.

4.3.1 Borrower Opacity

I perform several analyses that exploit cross-sectional variance in institutional lenders' information demand based on borrower characteristics in the pre-coverage period. First, I examine whether the effect of the satellite data coverage on institutional lending is more pronounced for opaque borrowers. Opaque borrowers provide imprecise public signals; therefore, traders have more heterogeneous beliefs about them, which encourages private information acquisition and informed trading (Verrecchia 1982; Diamond 1985; Bushman 1991). Moreover, when a borrower is opaque, its lenders have greater information advantages, which increases the value of the borrowers' private information. Therefore, institutional lenders should have a higher information demand for opaque borrowers when alternative information sources are not available.

To measure a borrower's information opacity, I begin with a borrower's analyst coverage. As an important information intermediary, financial analysts actively engage in private information production and provide accurate and timely information about firm performance to investors (Fried and Givoly 1982; Brown et al. 1987; Healy and Palepu 2001). Moreover, increased analyst following reduces the likelihood of insider trades and discourages insider purchases (Frankel and Li 2004). *No Analyst Coverage* equals 1 if the borrower does not have equity analyst coverage in the pre-coverage period (and 0 otherwise).

As another measure of borrower opacity, I consider a borrower's disclosure choices – decisions to issue earnings forecasts and press releases. Public disclosures may preclude costly private information acquisition (Diamond 1985; Verrecchia 2001) and are important determinants of a firm's information opacity (e.g., Beyer et al. 2010). *No Earnings Forecast* equals 1 if the

borrower does not issue earnings forecasts in the pre-coverage period (and 0 otherwise). *Low Press Releases* equals 1 if the borrower's average number of press releases in the pre-coverage period is less than the sample median (and 0 otherwise).

Using these borrower opacity variables, I assign loans to the high- and low-opacity partitions and re-estimate Model (1). In Panel A of Table 4, I find a negative and significant coefficient on *Treatment Firm* × *Post* in the low-analyst partition (*No Analyst Coverage* = 1). Importantly, I show that the magnitude of the coefficient on *Treatment Firm* × *Post* is statistically higher in the low-analyst partition than in the high-analyst partition. In Panel B of Table 4, the coefficients on *Treatment Firm* × *Post* are negative and significant using both the low-disclosure partition (*No Earnings Forecast* = 1) and high-disclosure partition (*No Earnings Forecast* = 0). However, the magnitude of the coefficient is statistically higher for the low-disclosure partition, consistent with non-guidance borrowers attracting higher information demand. In Panel C of Table 4, I find that the coefficient on *Treatment Firm* × *Post* is negative and significant in the low–press-release partition (*Low Press Releases* = 1) and its magnitude is statistically higher than the magnitude in the high–press-release partition (*Low Press Releases* = 0). Economically, using the low–press-release partition, the probability that an institutional lender issues a loan to a treatment borrower is 18.7% lower than for a control borrower in the coverage period. Overall, these results suggest that the satellite data coverage attenuates institutional lending to a greater extent for opaque borrowers.

4.3.2 Early Dissemination of Borrower Private Information

To strengthen the information demand mechanism, I perform additional cross-sectional tests to determine whether the satellite data coverage has a stronger effect when borrowers disseminate private information to their lenders earlier. Prior studies suggest that timely access to borrower information facilitates informed trading by incumbent lenders (Bushman et al. 2010; Carrizosa

and Ryan 2017). Because timely information is more valuable for instantaneous trading activities, I expect institutional lenders to have a higher information demand when borrowers disseminate their information to lenders on a timely basis.

I first measure early dissemination of borrower information based on whether a borrower issues a higher number of performance covenants (Bushman et al 2010, Christensen and Nikolaev 2012; Christensen et al. 2016; Carrizosa and Ryan 2017).²² Performance covenants are based on earnings and cash-flow metrics and they are frequently set tightly relative to the underlying performance variables. Moreover, these covenants often obligate borrowers to provide current performances information to lenders more frequently. Therefore, performance covenants enable lenders to monitor borrowers efficiently, which accelerates timely acquisition of private information about them (Bushman et al. 2010; Carrizosa and Ryan 2017). *High Perf. Covenants* equals 1 if the average number of performance covenants for loans issued to the borrower in the pre-coverage period is greater than the sample median (and 0 otherwise).

Next, I use lender reputation as another measure of timely dissemination of borrower information to lenders. The reputation of a lead arranger reflects its expertise and commitment to monitor borrowers (e.g., Diamond 1989; Boot et al. 1993; Chemmanur and Fulghieri 1994). Reputable lead arrangers collect greater private information about borrowers and communicate it earlier to syndicate participants (Bushman et al. 2010; Bushman and Wittenberg-Moerman 2012).²³ Therefore, I expect that, in the pre-coverage period, institutional lenders have a higher information demand when they participate in loans syndicated by reputable lead arrangers. *High*

²² Following Christensen and Nikolaev (2012), I classify cash interest coverage ratio, debt service coverage ratio, level of EBITDA, fixed charge coverage ratio, interest coverage ratio, ratio of debt to EBITDA, and ratio of senior debt to EBITDA covenants as performance covenants

²³ Also, reputable lenders incur higher reputational losses if they withhold important private information about borrowers from participants (Down et al. 2022).

Reputation equals 1 if the borrower obtains loans issued with one of the top five lead arrangers in the pre-coverage period (and 0 otherwise).

I partition sample observations based on these measures of timely dissemination of borrower information and re-estimate Model (1). As I report in Panel A of Table 5, the coefficient on *Treatment Firm* \times *Post* is significant in the high-covenant partition (*High Perf. Covenants* = 1) but not in the low-covenant partition (*High Perf. Covenants* = 0). I also show that the coefficient on *Treatment Firm* \times *Post* is statistically larger in the high-covenant partition than in the low-covenant partition. Further, Panel B of Table 5 shows a negative and significant coefficient on *Treatment Firm* \times *Post* in the high-reputation partition (*High Reputation* = 1) but not in the low-reputation partition (*High Reputation* = 0). In addition, the coefficient on *Treatment Firm* \times *Post* has a significantly higher magnitude in the high-reputation partition. Economically, using the high-reputation partition, the probability of an institutional lender issuing a loan to a treatment borrower in the coverage period is 29.4% lower than for a control borrower. Taken together, these findings suggest that the effect of satellite data coverage on institutional lending is greater when the flow of borrowers' private information to lenders is faster, which further supports the institutional lenders' information demand channel.

4.4 Accuracy of Satellite Image Data

In this section, I investigate whether the satellite data coverage has a greater effect on institutional lending when the data is more accurate in predicting borrowers' performance. When alternative sources of information provide signals with higher precision, traders can generate higher profits from informed trading (Grossman and Stiglitz 1980; McNichols and Trueman 1994). More precise satellite data can therefore further crowd out the value of private information acquired through lending relationships, leading to the reduction in institutional lender participation.

I measure the accuracy of the satellite data for each borrower using its store-level car counts.

A borrower has more accurate satellite data when the correlation between its car-count signals and firm performance is higher or when the variability of its car-count signals across stores is lower. *Treatment Firm High Corr* (*Treatment Firm Low Corr*) equals 1 if the average correlation between quarterly changes in the borrower’s car counts and its sales is greater (lower) than the sample median (and 0 otherwise). *Treatment Firm High SD* (*Treatment Firm Low SD*) equals 1 if the average standard deviation of quarterly changes in car counts across stores is greater (lower) than the sample median (and zero otherwise). Using each of these car-count-accuracy variables, I estimate the following model:

$$Inst. Lender = \beta_0 + \beta_1 Treatment Firm High Accuracy \times Post + \beta_2 Treatment Firm Low Accuracy \times Post + Controls + Fixed Effects + \varepsilon, \quad (3)$$

In this model, I replace *Treatment Firm* \times *Post* in Model (1) with separate interactions between *Post* and high (or low) accuracy of car-count variables.²⁴ In Panel A of Table 6, I present results of the analysis using *Treatment Firm High Corr* (and *Treatment Firm Low Corr*). I find a negative and significant coefficient on *Treatment Firm High Corr* \times *Post* but do not find a significant coefficient on *Treatment Firm Low Corr* \times *Post*; the difference between these two coefficients is statistically significant. I find similar results using *Treatment Firm High SD* (and *Treatment Firm Low SD*). In Panel B of Table 6, the coefficient on *Treatment Firm Low SD* \times *Post* is negative and significant across all specifications, and its magnitude is significantly higher than that of the coefficient on *Treatment Firm High SD* in OLS specifications. Economically, when the car-count signal exhibits lower variability, an institutional lender is 18.4% less likely to issue a loan to a treatment borrower in the coverage period. Overall, these results suggest that, when more precise

²⁴ Accuracy of the car-count signals cannot be measured for control borrowers whose car count data don’t exist.

satellite data further crowds out the value of borrowers' private information, institutional lenders have a lower demand to acquire such information by extending loans to borrowers. These findings not only further support the information demand mechanism but also provide evidence for the validity of the satellite data coverage as a proxy for the value of borrowers' private information.

4.5 Institutional Lenders' Demand for Private Information and Borrowing Terms

Thus far, I provide robust evidence that the information demand for borrowers' private information is an important factor for institutional lenders' decisions to issue loans. I next explore whether institutional lenders' information demand influences borrowers' credit outcomes. When the satellite data coverage reduces credit supply from institutional lenders, borrowers may obtain unfavorable loan terms (e.g., Ivashina and Sun 2011a; Lim et al. 2014). On the other hand, the satellite data provides useful information about borrower performance, which may mitigate syndicate participants' adverse selection concerns and incentivize them to supply more credit (Bushman et al. 2016; Kang et al. 2021b). In this case, borrowers may obtain favorable loan terms in the coverage period, despite institutional lenders' lower information demand. To investigate this question, I estimate the following OLS model:

$$\begin{aligned}
 \text{Loan Term} = & \beta_0 + \beta_1 \text{Treatment Firm} \times \text{Post No Inst. Lender} \times \text{Had Inst. Lender} + \\
 & \beta_2 \text{Treatment Firm} \times \text{Post Inst. Lender} \times \text{Had Inst. Lender} + \text{Main Effects} + \\
 & \text{Lower Order Interactions} + \text{Controls} + \text{Fixed Effects} + \varepsilon,
 \end{aligned}
 \tag{4}$$

where the dependent variable *Loan Term* is one of the following three borrowing terms: the natural logarithm of the all-in-drawn spread (*Interest Spread*), the natural logarithm of loan

amounts (*Amounts*), or the natural logarithm of the loan maturity in months (*Maturity*).²⁵ To estimate the effect of the reduction in institutional lenders' participation, I identify the following loans issued in the coverage period: loans issued without institutional lender participation (*Post No Inst. Lender*), loans issued with institutional lender participation (*Post Inst. Lender*), and loans issued to a borrower who had institutional lender participation in its loans issued in the pre-coverage period (*Had Inst. Lender*). The main variable of interest is the triple interaction term, *Treatment Firm* × *Post No Inst. Lender* × *Had Inst. Lender*. This variable measures loans issued to treatment borrowers (*Treatment Firm* = 1) who do not obtain loans from institutional lenders in the coverage period (*Post No Inst. Lender* = 1) but had lending relationships with institutional lenders in the pre-coverage period (*Had Inst. Lender* = 1), which indicates that these borrowers experienced a reduction in information demand from institutional lenders.

As I report in Table 7, I find a positive (negative) and significant coefficient on *Treatment Firm* × *Post No Inst. Lender* × *Had Inst. Lender* in the *Interest Spread* (*Amounts* or *Maturity*) specification. The results indicate that when institutional lenders stop issuing loans to borrowers in the coverage period, these borrowers pay higher interest rates, receive smaller loan amounts, and obtain loans with shorter maturities.²⁶ These unfavorable loan terms are consistent with reduced information demand leading to lower credit supply from institutional lenders.²⁷ Overall,

²⁵ I include all main effects and lower-order interactions of each triple interaction variable in Model (4) but do not specify them for brevity.

²⁶ I re-estimate Model (4) after restricting the sample to the treatment borrowers and continue to find a positive (negative) and significant coefficient on *Post No Inst. Lender* × *Had Inst. Lender* in *Interest Spread* (*Amounts* or *Maturity*) specification (reported in Panel F of Appendix B).

²⁷ I interpret the findings from Table 7 with caution because *Post No Inst. Lender* and *Had Inst. Lender* can be endogenously determined reflecting the institutional lenders' perceived costs and benefits from their lending relationships. In Panel G of Appendix B, I find that risk profiles of borrowers do not change significantly regardless of whether institutional lenders decide to sever or maintain their lending relationships in the coverage period. While these findings provide some reassurance, I acknowledge that there may be other unobserved factors that can potentially influence the decision of institutional lenders.

these findings suggest that institutional lender demand for borrowers' private information is an important factor that can influence contract outcomes of syndicated loans.

4.6. Satellite Image Data and Alternative Sources of Capital

The findings in Table 7 indicate that borrowers obtain unfavorable loan terms when institutional lenders' information demand decreases in the coverage period. This result raises an important question of whether the enhanced transparency brought by the availability of the satellite data can counterbalance the unfavorable loan terms by decreasing costs of capital from other sources. Specifically, the coverage of the satellite data may enhance the informativeness of stock prices, thereby reducing information asymmetry between firms and investors (Zhu 2019; Dichev and Qian 2022; Li and Venkatachalam 2022). The reduction in information asymmetry could subsequently lower the cost of raising equity capital (Easley and O'Hara 2004; Hughes et al. 2007). The availability of alternative data can also limit the propensity of insiders to trade based on private information, which can further contribute to a decrease in the cost of capital. On the other hand, the introduction of the satellite data coverage may increase the cost of capital because the data is accessible only to a subset of sophisticated investors. The unequal access to information can exacerbate information asymmetry between the sophisticated investors and individual investors, potentially raising cost of capital. Moreover, intensified short-selling following the initiation of the satellite data coverage can also make it costly to raise capital (Katona et al. 2021). Thus, the effect of the satellite data coverage on the cost of capital is not obvious.

To investigate whether the initiation of the satellite data coverage affects a borrower's cost of capital, I employ five commonly used measures of cost of capital: GLS (Gebhardt et al. 2001), CAT (Claus and Thomas 2001), PEG (Easton 2004), AGR (Ohlson and Juettner-Nauroth 2005), and AVG—the equally-weighted average of GLS, CAT, PEG, and AGR proxies (Lee et al. 2021). Using these measures of cost of capital as dependent variables, I re-estimate Model (1) and report

the results in Panel A of Table 8. The results do not indicate a significant impact of the initiation of the satellite data coverage on a borrower's cost of capital. The coefficients on *Treatment Firm x Post* are not statistically significant across analyses using each of GLS, CAT, PEG, and AGR proxies. However, using AVG, I find a positive and significant coefficient on *Treatment Firm x Post*. These results provide weak evidence suggesting that a borrower's cost of capital may increase in the coverage period. I employ AVG as a proxy for cost of capital and further examine the effect of the satellite data coverage for opaque borrowers who induce higher information demand from institutional lenders. Using *No Analyst Coverage*, *No Earnings Forecast*, and *Low Press Releases* as measures of opaque borrowers, I re-estimate the model after including *Treatment Firm x Post x Opaque Borrowers* and *Post x Opaque Borrowers*.²⁸ As reported in Panel B of Table 8, the sum of coefficients on *Treatment Firm x Post x Opaque Borrowers* and *Treatment Firm x Post* is not statistically different from zero across analyses.²⁹ The results indicate that cost of capital does not change for opaque borrowers in the coverage period.

To examine whether borrowers substitute for the reduced credit supply from institutional lenders by issuing more equity, I collect data on new equity issuance from SDC and construct the following two measures of borrowers' equity issuance: *Equity Issuance Amount* is the natural logarithm of the equity amounts raised by the borrower, and *Equity Issuance Indicator* is an indicator variable equal to 1 if the borrower issues new equity, and 0 otherwise. I re-estimate Model (1) using these measures as dependent variables and report the results in Panel C of Table 8. I find that the coefficients on *Treatment Firm x Post* are insignificant across analyses, which

²⁸ Other lower-order interaction terms and main effects are absorbed by firm and year-quarter fixed effects.

²⁹ The negative coefficients on *Treatment Firm x Post x Opaque Borrowers* indicate that cost of capital is lower for opaque borrowers relative to non-opaque borrowers in the coverage period, which suggests that the enhanced transparency after the initiation of the satellite data coverage may disproportionately benefit opaque borrowers.

suggests that borrowers do not issue equity in the coverage period. I also investigate whether opaque borrowers are more likely to issue equity in the coverage period by re-estimating the model after including *Treatment Firm x Post x Opaque Borrowers* and *Post x Opaque Borrowers*. In the Panel D of Table 8, I find that both coefficients on *Treatment Firm x Post* and *Treatment Firm x Post x Opaque Borrowers* are insignificant across analyses, which indicates that opaque borrowers also do not issue additional equity in the coverage period. These results are consistent with the previous finding that the availability of the satellite data does not translate into reduced cost of capital.

Although I do not find evidence of borrowers using alternative sources of financing to compensate for the loss of credit supply from institutional lenders in the coverage period, I interpret these results with caution. These results do not directly address the overall welfare implications of the satellite data coverage for borrowers. The impact of the satellite data coverage on borrowers is likely multifaceted and may be influenced by a range of factors not captured in these analyses. Further research is needed to fully understand the broader implications of the satellite data coverage on borrowers' financial condition and overall welfare.

4.7. Institutional Lenders' Information Demand and Pricing Dynamics in the Institutional Loan Market

In the final set of analyses, I explore whether institutional lenders' information demand, particularly for opaque borrowers, affects loan spreads in the coverage period.³⁰ Prior studies in compensation literature suggest that the optimal contract reduces explicit wages paid to managers when there are insider trading opportunities because the managers are also remunerated through the expected insider trading profits (Baiman and Verrecchia 1996;

³⁰ In Section A of the online appendix, I find some evidence that opaque borrowers are more likely to obtain loans from institutional lenders in the pre-coverage period. This result further demonstrates that the information demand can influence institutional lenders' lending decisions.

Roulstone 2003). Applying these insights to the institutional loan market, total return to institutional lenders can be conceptualized as explicit interest spreads on the loan plus insider trading profits. Therefore, institutional lenders may accept lower interest spreads if they expect significant insider trading profits. This leads to a prediction that, in the pre-coverage period, institutional lenders' information demand puts downward pressure on interest spreads for loans to opaque borrowers who present greater opportunities for insider trading.

To test this prediction, I restrict the sample to the pre-coverage period and estimate the following OLS model:

$$Interest\ Spread = \beta_0 + \beta_1 Inst.\ Lender + \beta_2 Borrower\ Opacity + \beta_3 Inst.\ Lender \times Borrower\ Opacity + Controls + Fixed\ Effects + \varepsilon, \quad (5)$$

where *Borrower Opacity* is one of the following variables to capture opaque borrowers who (a) lack analyst coverage (*No Analyst Coverage*), (b) do not issue earnings forecasts (*No Earnings Forecast*), or (c) issue fewer press releases (*Low Press Releases*). I control for borrower characteristics included in Model (1) and include year–quarter, credit-rating and loan-type fixed effects.³¹

The results of the analyses, presented in Table 9, do not support the prediction that institutional lenders reduce interest spreads when they issue loans to opaque borrowers in the pre-coverage period. In fact, the coefficient on *Inst. Lender × No Analyst Coverage* is positive and significant. This suggests that, contrary to the prediction, interest spreads are higher for institutional loans issued to opaque borrowers who lack analyst coverage.³² This finding is consistent with prior studies indicating that institutional lenders have greater bargaining power

³¹ I exclude firm fixed effects because *Borrower Opacity* is measured at the firm level.

³² The coefficients on *Inst. Lender × No Earnings Forecast* and *Inst. Lender × Low Press Releases* are not statistically significant, which suggests that interest spreads remain unchanged for institutional loans issued to borrowers who do not issue earnings forecasts or who issue fewer press releases.

because they often serve as lenders of last resort and primarily issue loans to riskier borrowers, leading to higher interest spreads (Taylor and Sansone 2006; Nandy and Shao 2008; Ivashina and Sun 2011b; Lim et al. 2014). Therefore, although the information demand puts downward pressure on interest spreads, institutional lenders' bargaining power counteracts this effect so that they do not have to lower interest spreads.³³ These results highlight the intricate interplay of various factors that can shape the complex pricing dynamics within the institutional loan market.

5. Conclusion

I show that the value of borrowers' private information is a significant determinant for institutional lenders' participation in syndicated loans. As a shock to institutional lenders' private information advantages, I utilize the release of the satellite image data of car counts in store parking lots of U.S. retail firms. I predict that accurate and near-real-time information on borrower performance through the satellite data diminishes the value of borrowers' private information; therefore, institutional lenders have a lower demand for the private information obtained through lending relationships. Consistent with my prediction, I find that institutional lenders are less likely to participate in loan syndicates after the satellite data on a borrower becomes commercially available. Supporting the information demand argument, I further show that the satellite data coverage further attenuates institutional lending when borrowers are opaque, when they disseminate private information to their lenders earlier, or when the satellite data provides more accurate forecasts of borrower performance. Lastly, I find that institutional lenders' lower information demand leads to unfavorable credit outcomes for borrowers in the

³³ In Section B of the online appendix, I do not find evidence that institutional lenders reduce interest spreads for opaque borrowers even in competitive credit market environments in which (a) net inflow of funds into the institutional loan market is higher, (b) the average number of days that syndicated loans remain unsold after the launch day is smaller, and (c) the percentage of banks tightening standards for commercial and industrial loans to large and middle-market firms is lower.

coverage period, while greater information demand in the pre-coverage period—especially for opaque borrowers—does not result in lower premium.

My study is not without limitations. My sample is restricted to retail firms because satellite images of store parking lots are available only for those firms. Although I believe that institutional investors' information demand is an important determinant of their participation in loan syndicates, I caution against generalizing my results to firms in other industries. I leave it for future research to explore whether the information demand significantly influences institutional lending and credit outcomes for non-retailer borrowers. In addition, future research can also identify other sources of big data and examine how institutional lenders' information demand varies with unique features of these data.

REFERENCES

- Acharya, V., and T. Johnson. 2007. Insider trading in credit derivatives. *Journal of Financial Economics*. 84 (1), 110-141.
- Akins, B., J. Ng, and R. Verdi. 2012. Investor competition over information and the pricing of information asymmetry. *The Accounting Review*. 87 (1), 35-58.
- Amiraslani, H., J. Donovan, M. Phillips, and R. Wittenberg-Moerman. 2023. Contracting in the Dark: The rise of public-side lenders in the syndicated loan market. *Journal of Accounting and Economics*. 101586.
- Angrist, J., and J. Pischke. 2008. Mostly harmless econometrics: An empiricist's companion: Princeton University Press.
- Back, K., C. Cao, and G. Willard. 2000. Imperfect competition among informed traders. *The Journal of Finance*. 55 (5), 2117-2155.
- Bartlett, R., A. Morse, R. Stanton and N. Wallace 2022. Consumer-lending discrimination in the FinTech era. *Journal of Financial Economics*. 143(1), 30-56.
- Beyer, A., D. Cohen, T. Lys, and B. Walther. 2010. The financial reporting environment: Review of the recent literature. *Journal of Accounting and Economics*. 50 (2-3), 296-343.
- Baiman, S., and R. E. Verrecchia. 1996. The relation among capital markets, financial disclosure, production efficiency, and insider trading. *Journal of Accounting Research*. 34 (1), 1-22.
- Blankespoor, E., E. deHaan, and I. Marinovic. 2020. Disclosure processing costs, investors' information choice, and equity market outcomes: A review. *Journal of Accounting and Economics*. 70 (2-3), 101344.
- Boot, A., S. Greenbaum, and A. Thakor. 1993. Reputation and discretion in financial contracting. *The American Economic Review*. 1165-1183.
- Brown, L., R. Hagerman, P. Griffin, and M. Zmijewski. 1987. Security analyst superiority relative to univariate time-series models in forecasting quarterly earnings. *Journal of Accounting and Economics*. 9 (1), 61-87.
- Bushman, R. 1991. Public disclosure and the structure of private information markets. *Journal of Accounting Research*. 29 (2), 261-276.
- Bushman, R., A. J. Smith, and R. Wittenberg-Moerman. 2010. Price discovery and dissemination of private information by loan syndicate participants. *Journal of Accounting Research*. 48 (5), 921-972.
- Bushman, R., and R. Wittenberg-Moerman. 2012. The role of bank reputation in "certifying" future performance implications of borrowers' accounting numbers. *Journal of Accounting Research*. 50 (4), 883-930.
- Bushman, R., C. Williams, and R. Wittenberg-Moerman. 2016. The Informational Role of the Media in Private Lending. *Journal of Accounting Research*. 55 (1), 115-152.
- Carey, M., M. Post, and S. Sharpe. 1998. Does corporate lending by banks and finance companies differ? Evidence on specialization in private debt contracting. *The Journal of Finance*. 53 (3), 845-878.

- Carrizosa, R., and S. Ryan. 2017. Borrower private information covenants and loan contract monitoring. *Journal of Accounting and Economics*. 64 (2-3), 313-339.
- Chen, W., J. Kang and A. Mohan. 2023. Data-Driven Technologies and the Local Information Advantages in Small Business Lending. Working Paper.
- Chemmanur, T., and P. Fulghieri. 1994. Investment bank reputation, information production, and financial intermediation. *The Journal of Finance*. 49 (1), 57-79.
- Chernenko, S., I. Erel, and R. Prilmeier. 2019. Nonbank lending. Working paper.
- Christensen, H., and V. Nikolaev. 2012. Capital versus performance covenants in debt contracts. *Journal of Accounting Research*. 50 (1), 75-116.
- Christensen, H., V. Nikolaev, and R. Wittenberg-Moerman. 2016. Accounting information in financial contracting: The incomplete contract theory perspective. *Journal of Accounting Research*. 54 (2), 397-435.
- Claus, J., and J. Thomas. 2001. Equity premia as low as three percent? Evidence from analysts' earnings forecasts for domestic and international stock markets. *The Journal of Finance*. 56 (5), 1629-1666.
- Costello, A., A. Down and M. Mehta. 2020. Machine+ man: A field experiment on the role of discretion in augmenting AI-based lending models. *Journal of Accounting and Economics*. 70(2-3), 101360.
- Down, A., C. Williams, and R. Wittenberg-Moerman. 2022. Strategic syndication: is bad news shared in loan syndicates? *Review of Accounting Studies*. 1-43.
- Diamond, D. 1985. Optimal release of information by firms. *The Journal of Finance*. 40 (4), 1071-1094.
- Diamond, D. 1989. Reputation acquisition in debt markets. *Journal of Political Economy*. 97 (4), 828-862.
- Dichev, I. D., and J. Qian. 2022. The benefits of transaction-level data: The case of NielsenIQ scanner data. *Journal of Accounting and Economics*. 74 (1), 101495.
- Drucker, S., and M. Puri. 2005. On the benefits of concurrent lending and underwriting. *The Journal of Finance*. 60 (6), 2763-2799.
- Easley, D., and M. O'Hara. 2004. Information and the cost of capital. *The Journal of Finance*. 59 (4), 1553-1583.
- Easton, P. D. 2004. PE ratios, PEG ratios, and estimating the implied expected rate of return on equity capital. *The Accounting Review*. 79 (1), 73-95.
- FDIC. 2019. Bank and nonbank lending over the past 70 years. *FDIC Quarterly*. 31-39
- Foster, F., and S. Viswanathan. 1993. The effect of public information and competition on trading volume and price volatility. *The Review of Financial Studies*. 6 (1), 23-56.
- Foster, F., and S. Viswanathan. 1996. Strategic trading when agents forecast the forecasts of others. *The Journal of Finance*. 51 (4), 1437-1478.

- Frankel, R., and X. Li. 2004. Characteristics of a firm's information environment and the information asymmetry between insiders and outsiders. *Journal of Accounting and Economics*. 37 (2), 229-259.
- Fried, D., and D. Givoly. 1982. Financial analysts' forecasts of earnings: A better surrogate for market expectations. *Journal of Accounting and Economics*. 4 (2), 85-107.
- Gebhardt, W. R., C. M. Lee, and B. Swaminathan. 2001. Toward an implied cost of capital. *Journal of Accounting Research*. 39 (1), 135-176.
- Glaeser, S., B. Schonberger, C. Wasley, and J. Xiao. 2023. Private Information Acquisition via Freedom of Information Act Requests Made to the Securities and Exchange Commission. *The Accounting Review*. 98 (3), 229-255.
- Greene, W. (2004). The behaviour of the maximum likelihood estimator of limited dependent variable models in the presence of fixed effects. *The Econometrics Journal*. 7 (1), 98-119.
- Grossman, S., and J. Stiglitz. 1980. On the impossibility of informationally efficient markets. *The American Economic Review*. 70 (3), 393-408.
- Gustafson, M., I. Ivanov, and R. Meisenzahl. 2021. Bank monitoring: Evidence from syndicated loans. *Journal of Financial Economics* 139 (2), 452-477.
- Hainmueller, J. 2012. Entropy balancing for causal effects: A multivariate reweighting method to produce balanced samples in observational studies. *Political Analysis*. 20 (1), 25-46.
- Han, S., and X. Zhou. 2014. Informed bond trading, corporate yield spreads, and corporate default prediction. *Management Science*. 60 (3), 675-694.
- Haselmann, R., C. Leuz, and S. Schreiber. 2022. Know your customer: Informed trading by banks. Working Paper.
- Healy, P., and K. Palepu. 2001. Information asymmetry, corporate disclosure, and the capital markets: A review of the empirical disclosure literature. *Journal of Accounting and Economics*. 31 (1-3), 405-440.
- Holden, C., and A. Subrahmanyam. 1992. Long-lived private information and imperfect competition. *The Journal of Finance*. 47 (1), 247-270.
- Hughes, J. S., J. Liu, and J. Liu. 2007. Information asymmetry, diversification, and cost of capital. *The Accounting Review*. 82 (3), 705-729.
- Irani, R., R. Iyer, R. Meisenzahl, and J.-Peydró. 2021. The rise of shadow banking: evidence from capital regulation. *The Review of Financial Studies*. 34 (5), 2181-2235.
- Ivashina, V., and Z. Sun. 2011a. Institutional demand pressure and the cost of corporate loans. *Journal of Financial Economics*. 99 (3), 500-522.
- Ivashina, V., and Z. Sun. 2011b. Institutional stock trading on loan market information. *Journal of Financial Economics*. 100 (2), 284-303.
- Jame, R., R. Johnston, S. Markov, and M. Wolfe. 2016. The value of crowdsourced earnings forecasts. *Journal of Accounting Research*. 54 (4), 1077-1110.

- Jiang, W., K. Li, and P. Shao. 2010. When shareholders are creditors: Effects of the simultaneous holding of equity and debt by non-commercial banking institutions. *The Review of Financial Studies*. 23 (10), 3595-3637.
- Kang, J., L. Stice-Lawrence, and Y. Wong. 2021a. The firm next door: Using satellite images to study local information advantage. *Journal of Accounting Research*. 59 (2), 713-750.
- Kang, J., C. Williams, and R. Wittenberg-Moerman. 2021b. CDS trading and nonrelationship lending dynamics. *Review of Accounting Studies*. 26 (1), 258-292.
- Kang, J., M. Loumioti and R. Wittenberg-Moerman 2021c. The harmonization of lending standards within banks through mandated loan-level transparency. *Journal of Accounting and Economics*. 72(1), 101386.
- Kang, J., C. Lennox, and V. Pandey. 2022. Client concerns about information spillovers from sharing audit partners. *Journal of Accounting and Economics*. 73 (1), 101434.
- Katona, Z., M. Painter, P. Patatoukas, and J. Zeng. 2021. On the capital market consequences of alternative data: Evidence from outer space. Working Paper.
- Kim, O., and R. Verrecchia. 1991. Trading volume and price reactions to public announcements. *Journal of Accounting Research*. 29 (2), 302-321.
- Kim, O., and R. E. Verrecchia. 1994. Market liquidity and volume around earnings announcements. *Journal of Accounting and Economics*. 17 (1-2), 41-67.
- Kyle, A. 1985. Continuous auctions and insider trading. *Econometrica: Journal of the Econometric Society*. 1315-1335.
- Lee, C. M., E. C. So, and C. C. Wang. 2021. Evaluating firm-level expected-return proxies: implications for estimating treatment effects. *The Review of Financial Studies*. 34 (4), 1907-1951.
- Li, B., and M. Venkatachalam. 2022. Leveraging big data to study information dissemination of material firm events. *Journal of Accounting Research*. 60 (2), 565-606.
- Lim, J., B. Minton, and M. Weisbach. 2014. Syndicated loan spreads and the `ion of the syndicate. *Journal of Financial Economics*. 111 (1), 45-69.
- Loumioti, M. 2022. Direct Lending: The determinants, characteristics and performance of direct loans. Working Paper.
- Liu, M. 2022. Assessing human information processing in lending decisions: A machine learning approach. *Journal of Accounting Research*. 60 (2), 607-651.
- Maddala, G. 1987. Limited dependent variable models using panel data. *Journal of Human Resources*. 22 (3), 307-338.
- Massoud, N., D. Nandy, A. Saunders, and K. Song. 2011. Do hedge funds trade on private information? Evidence from syndicated lending and short-selling. *Journal of Financial Economics*. 99 (3), 477-499.
- McMullin, J., and B. Schonberger. 2020. Entropy-balanced accruals. *Review of Accounting Studies*. 25 (1), 84-119

- McNichols, M., and B. Trueman. 1994. Public disclosure, private information collection, and short-term trading. *Journal of Accounting and Economics*. 17 (1-2), 69-94.
- Minnis, M., A. Sutherland, and F. Vetter. 2023. Financial statements not required. Working Paper.
- Nandy, D. K., and P. Shao. 2008. Institutional investment in syndicated loans. Working Paper.
- Ohlson, J. A., and B. E. Juettner-Nauroth. 2005. Expected EPS and EPS growth as determinants of value. *Review of Accounting Studies*. 10, 349-365.
- Peyravan, L. 2020. Financial reporting quality and dual-holding of debt and equity. *The Accounting Review*. 95 (5), 351-371.
- Peyravan, L., and R. Wittenberg-Moerman. 2022. Institutional dual-holders and managers' earnings disclosure. *The Accounting Review*. 97 (3), 343-371.
- Roulstone, D. T. 2003. The relation between insider-trading restrictions and executive compensation. *Journal of Accounting Research*. 41 (3), 525-551.
- Securities and Exchange Commission (SEC). 2012. Summary Report on Examinations of Information Barriers: Broker-Dealer Practices Under Section 15(g) of the Securities Exchange Act of 1934. Office of Compliance Inspections and Examinations. Washington DC: SEC.
- Standard & Poor's. 2020. Leveraged Commentary & Data (LCD): Leveraged loan primer. Global Market Intelligence Inc.
- Sutherland, A. 2018. Does credit reporting lead to a decline in relationship lending? Evidence from information sharing technology. *Journal of Accounting and Economics*. 66(1), 123-141.
- Taylor, A., and A. Sansone. 2006. The handbook of loan syndications and trading: McGraw Hill Professional.
- Verrecchia, R. 1982. Information acquisition in a noisy rational expectations economy. *Econometrica: Journal of the Econometric Society*. 1415-1430.
- Verrecchia, R. 2001. Essays on disclosure. *Journal of Accounting and Economics*. 32 (1-3), 97-180.
- Yasuda, A. 2005. Do bank relationships affect the firm's underwriter choice in the corporate-bond underwriting market? *The Journal of Finance*. 60 (3), 1259-1292.
- Zhu, C. 2019. Big data as a governance mechanism. *The Review of Financial Studies*. 32 (5), 2021-2061.

APPENDIX A
Variable Definitions

Variable	Definition
<i>Age</i>	= The number of years since a firm first appears in Compustat (Compustat).
<i>Altman Z</i>	= Altman (1963) Z-score as estimated by the following model: $Z = 3.3X_1 + 0.99X_2 + 0.6X_3 + 1.2X_4 + 1.4X_5$, where X_1 is the ratio of earnings before interest and taxes to total assets, X_2 is the ratio of total sales to total assets, X_3 is the ratio of market value of equity to total liabilities, X_4 is the ratio of current assets to total assets, and X_5 is the ratio of retained earnings to total assets. All variables are measured in the year preceding the loan's issuance (Compustat).
<i>Amounts</i>	= The natural logarithm of loan amounts of the largest facility in the loan package (DealScan).
<i>Assets</i>	= The natural logarithm of total assets, measured in the year preceding the loan's issuance (Compustat).
<i>Current Ratio</i>	= The ratio of current assets to current liabilities, measured in the year preceding the loan's issuance (Compustat).
<i>Equity Issuance Amount</i>	= The natural logarithm of equity amounts raised by the borrower, measured in the year of the loan's issuance (SDC).
<i>Equity Issuance Indicator</i>	= An indicator variable equal to 1 if the borrower issues equity in the year of the loan's issuance, and 0 otherwise (SDC).
<i>Guarantor</i>	= An indicator variable equal to 1 if the loan is guaranteed, and 0 otherwise (DealScan).
<i>Had Inst. Lender</i>	= An indicator variable equal to 1 if the loan is issued after the satellite image data becomes commercially available and the borrower had lending relationships with institutional lenders (<i>Inst. Lender</i>) before the satellite image data becomes commercially available, and 0 otherwise (Orbital Insight).
<i>High Perf. Covenants</i>	= An indicator variable equal to 1 if the average number of performance covenants for loans issued to the borrower, in the pre-coverage period before the satellite image data becomes commercially available, is greater than the sample median, and 0 otherwise (DealScan).
<i>High Reputation</i>	= An indicator variable equal to 1 if the borrower obtains loans issued with one of the top five lead arrangers in the pre-coverage period before the satellite image data becomes commercially available, and 0 otherwise (DealScan, Bloomberg).
<i>ICC</i>	= An internal rate of return that equates a firm's forecasted cash flows to its current market price: <i>AGR ICC</i> following Ohlson and Juettner-Nauroth (2005), <i>CAT ICC</i> following Claus and Thomas (2001), <i>GLS ICC</i> following Gebhardt et al. (2001), and <i>PEG ICC</i> following Easton (2004). <i>AVG ICC</i> is an equally-weighted average of the following four measures of cost of capital: <i>AGR</i> (Ohlson and Juettner-Nauroth 2005), <i>CAT</i> (Claus and Thomas 2001), <i>GLS</i> (Gebhardt et al. 2001), and <i>PEG</i> (Easton 2004).

APPENDIX A (continued)
Variable Definitions

Variable	Definition
<i>Inst. Lender Bank Affiliated</i>	= An indicator variable equal to 1 if the loan is issued with at least one bank-affiliated institutional lender but is not issued with an independent institutional lender, and 0 otherwise (DealScan).
<i>Inst. Lender</i>	= An indicator variable equal to 1 if the loan is issued with at least one independent institutional lender that engages in investment businesses, and 0 otherwise (DealScan).
<i>Inst. Lender No Investment</i>	= An indicator variable equal to 1 if the loan is issued with at least one independent institutional lender that does <i>not</i> engage in investment businesses but is not issued with an independent institutional lender that does engage in investment businesses, and 0 otherwise (DealScan).
<i>Interest Coverage</i>	= The ratio of earnings before interest and taxes to interest expense, measured in the year preceding the loan's issuance (Compustat).
<i>Interest Spread</i>	= The natural logarithm of the all-in-drawn spread of the largest facility in the package (DealScan).
<i>Leverage</i>	= The ratio of long-term debt plus debt in current liabilities to total assets, measured in the year preceding the loan's issuance (Compustat).
<i>Low Press Releases</i>	= An indicator variable equal to 1 if the average number of press releases by the borrower, measured in the pre-coverage period before the satellite image data becomes commercially available, is less than the sample median, and 0 otherwise (RavenPack).
<i>Maturity</i>	= The natural logarithm of the loan maturity in months (DealScan).
<i>MTB</i>	= The ratio of market value to book value of equity, measured in the year preceding the loan's issuance (Compustat).
<i>No Analyst Coverage</i>	= An indicator variable equal to 1 if the borrower does not have equity analyst coverage, measured in the pre-coverage period before the satellite image data becomes commercially available, and 0 otherwise (IBES).
<i>No Earnings Forecast</i>	= An indicator variable equal to 1 if the borrower does not issue earnings forecasts, measured in the pre-coverage period before the satellite image data becomes commercially available, and 0 otherwise (IBES).
<i>Past Return</i>	= The accumulated daily stock return measured over 150 calendar days ending 30 days before the loan's issuance (DealScan, CRSP).
<i>Post</i>	= An indicator variable equal to 1 if the loan is issued after the satellite image data becomes commercially available, and 0 otherwise (Orbital Insight).
<i>Post Inst. Lender</i>	= An indicator variable equal to 1 if the loan is issued with institutional lenders (<i>Inst. Lender</i>) after the satellite image data becomes commercially available, and 0 otherwise (Orbital Insight).

APPENDIX A (continued)
Variable Definitions

Variable	Definition
<i>Post No Inst. Lender</i>	= An indicator variable equal to 1 if the loan is not issued with institutional lenders (<i>Inst. Lender</i>) after the satellite image data becomes commercially available, and 0 otherwise (Orbital Insight).
<i>ROA</i>	= The ratio of net income to total assets, measured in the year preceding the loan's issuance (Compustat).
<i>Sales Growth</i>	= The ratio of total sales in year t to total sales in year $t-1$ minus 1, measured in the year preceding the loan's issuance (Compustat).
<i>Secured</i>	= An indicator variable equal to 1 if the loan is secured, and 0 otherwise (DealScan).
<i>Treatment Firm</i>	= An indicator variable equal to 1 if the borrower is covered by the satellite data after it becomes commercially available (Orbital Insight).
<i>Treatment Firm High Corr</i>	= An indicator variable equal to 1 if the average correlation between quarterly changes in store-level car counts and quarterly changes in the borrower's sales, measured in the coverage period after the satellite image data becomes commercially available, is greater than the sample median, and 0 otherwise (Orbital Insight).
<i>Treatment Firm Low Corr</i>	= An indicator variable equal to 1 if the average correlation between quarterly changes in store-level car counts and quarterly changes in the borrower's sales, measured in the coverage period after the satellite image data becomes commercially available, is less than the sample median, and 0 otherwise (Orbital Insight).
<i>Treatment Firm High SD</i>	= An indicator variable equal to 1 if the average standard deviation of quarterly changes in car counts across stores, measured in the post-period after the satellite image data becomes commercially available, is greater than the sample median, and 0 otherwise (Orbital Insight).
<i>Treatment Firm Low SD</i>	= An indicator variable equal to 1 if the average standard deviation of quarterly changes in car counts across stores, measured in the post-period after the satellite image data becomes commercially available, is less than the sample median, and 0 otherwise (Orbital Insight).

APPENDIX B
Additional Analyses

This table reports the results of additional analyses. Panel A examines whether the main results are robust to using continuous dependent variables that capture the extent of institutional lender participation. Column 1 estimates a Tobit model using the dependent variable *Inst. Lender Proportion*, which is the proportion (%) of institutional lenders in the loan package. Column 2 estimates a Poisson model using the dependent variable *Inst. Lender Counts*, which is the number of institutional lenders in the loan package. Panel B reports the result of analysis whether the main results are robust to using *Post RM*, an indicator variable equal to 1 if the loan is issued after the satellite image data from either RS Metrics or Orbital Insight becomes commercially available, and 0 otherwise. Panel C tests whether the main results are robust to excluding sample periods when SafeGraph released their mobile GPS location data in 2018. Panel D restricts the main sample to the pre-coverage period and re-estimates Model (1) after interacting *Treatment Firm* with a continuous trend variable of *Year*, which is the year of loan issuance. Panel E compares the mean and standard deviation of the explanatory variables for the treatment and control firms to provide evidence of covariate balancing in the estimation using an entropy balancing approach. Panel F examines, using the treatment sample, whether the reduction in institutional lenders' information demand affects borrowers' credit outcomes. Panel G examines whether there are changes in a borrower's risk profile when institutional lenders terminate the lending relationship after the satellite data on the borrower becomes commercially available. t-statistics in parentheses are based on standard errors clustered at the firm level. ***, **, and * indicate significance at the 0.01, 0.05, and 0.10 level, respectively. All other variables are defined in Appendix A.

Panel A: Continuous Dependent Variables

	<i>Inst. Lender Proportion</i>	<i>Inst. Lender Counts</i>
	(1)	(2)
<i>Treatment Firm x Post</i>	-25.589***	-0.772*
	(-3.05)	(-1.95)
<i>Assets</i>	0.415 (0.25)	-0.035 (-0.22)
<i>Current Ratio</i>	0.806 (0.40)	0.166 (1.22)
<i>Leverage</i>	-8.005 (-1.05)	0.033 (0.08)
<i>MTB</i>	0.287 (1.50)	0.007 (0.78)
<i>Sales Growth</i>	3.151 (0.68)	0.330 (1.57)
<i>Interest Coverage</i>	-0.015* (-1.85)	-0.001* (-1.82)
<i>ROA</i>	-45.120*** (-3.14)	-0.862 (-1.12)
<i>Altman Z</i>	-0.203 (-0.26)	-0.041 (-0.84)

APPENDIX B (continued)
Additional Analyses

Panel A: Continuous Dependent Variables (continued)

	<i>Inst. Lender Proportion</i>	<i>Inst. Lender Counts</i>
	(1)	(2)
<i>AGE</i>	-0.005 (-0.03)	-0.009 (-0.49)
<i>Past Return</i>	-9.600 (-1.56)	-0.351 (-1.17)
<i>Amounts</i>	1.664 (0.77)	0.109 (1.07)
<i>Maturity</i>	15.954* (1.91)	1.009** (2.12)
<i>Secured</i>	13.197*** (3.36)	0.483** (2.09)
<i>Guarantor</i>	3.138 (0.53)	0.345 (0.89)
<i>Model</i>	Tobit	Poisson
<i>Firm FE</i>	Yes	Yes
<i>Year FE</i>	Yes	Yes
<i>Year-Quarter FE</i>	No	No
<i>Credit Rating FE</i>	No	No
<i>Loan Type FE</i>	No	No
<i>Observations</i>	2,129	945

Panel B: RS Metrics Data

	<i>Inst. Lender</i>		
	(1)	(2)	(3)
<i>Post RM</i>	-0.928* (-1.77)	-0.090** (-2.14)	-0.120*** (-2.77)
<i>Model</i>	Logit	OLS	OLS
<i>Controls</i>	Yes	Yes	Yes
<i>Firm FE</i>	Yes	Yes	Yes
<i>Year FE</i>	Yes	Yes	No
<i>Year-Quarter FE</i>	No	No	Yes
<i>Credit Rating FE</i>	No	No	Yes
<i>Loan Type FE</i>	No	No	Yes
<i>Observations</i>	904	2,129	2,129
<i>Adj. (Pseudo) R-squared</i>	0.066	0.228	0.245

APPENDIX B (continued)
Additional Analyses

Panel C: Excluding Periods after 2018

	<i>Inst. Lender</i>		
	(1)	(2)	(3)
<i>Treatment Firm x Post</i>	-1.515** (-2.46)	-0.148*** (-3.02)	-0.168*** (-3.51)
<i>Model</i>	Logit	OLS	OLS
<i>Controls</i>	Yes	Yes	Yes
<i>Firm FE</i>	Yes	Yes	Yes
<i>Year FE</i>	Yes	Yes	No
<i>Year-Quarter FE</i>	No	No	Yes
<i>Credit Rating FE</i>	No	No	Yes
<i>Loan Type FE</i>	No	No	Yes
<i>Observations</i>	721	1,866	1,866
<i>Adj. (Pseudo) R-squared</i>	0.099	0.232	0.245

Panel D: Continuous Trend Variable

	<i>Inst. Lender</i>		
	(1)	(2)	(3)
<i>Treatment Firm x Year</i>	0.111 (0.55)	0.008 (0.37)	0.003 (0.12)
<i>Model</i>	Logit	OLS	OLS
<i>Controls</i>	Yes	Yes	Yes
<i>Firm FE</i>	Yes	Yes	Yes
<i>Year FE</i>	Yes	Yes	No
<i>Year-Quarter FE</i>	No	No	Yes
<i>Credit Rating FE</i>	No	No	Yes
<i>Loan Type FE</i>	No	No	Yes
<i>Observations</i>	446	1,422	1,422
<i>Adj. (Pseudo) R-squared</i>	0.118	0.214	0.215

APPENDIX B (continued)
Additional Analyses

Panel E: Covariate Balancing

	<i>Pre-Matching</i>					<i>Post-Matching</i>				
	Treatment Mean	Control Mean	Treatment SD	Control SD	Diff Mean	Treatment Mean	Control Mean	Treatment SD	Control SD	Diff Mean
<i>Assets</i>	8.174	8.038	1.549	1.839	0.136	8.174	8.173	1.549	1.549	0.001
<i>Current Ratio</i>	1.581	1.602	0.86	1.004	-0.021	1.581	1.58	0.86	0.86	0.001
<i>Leverage</i>	0.311	0.356	0.276	0.264	-0.045***	0.311	0.311	0.276	0.276	0.000
<i>MTB</i>	3.202	3.237	9.83	8.847	-0.035	3.202	3.202	9.83	9.827	0.000
<i>Sales Growth</i>	0.054	0.171	0.142	0.378	-0.117***	0.054	0.054	0.142	0.142	0.000
<i>Interest Coverage</i>	72.63	63.79	268.132	291.148	8.84	72.63	72.63	268.132	268.157	0.000
<i>ROA</i>	0.062	0.019	0.077	0.127	0.043***	0.062	0.062	0.077	0.077	0.000
<i>Altman Z</i>	4.673	3.276	2.277	2.909	1.397***	4.673	4.672	2.277	2.277	0.001
<i>AGE</i>	28.96	23.03	17.433	18.73	5.93***	28.96	28.95	17.433	17.433	0.010
<i>Past Return</i>	0.056	0.055	0.245	0.253	0.001	0.056	0.056	0.245	0.245	0.000
<i>Amounts</i>	19.89	19.8	1.031	1.048	0.09	19.89	19.89	1.031	1.031	0.000
<i>Maturity</i>	3.997	3.95	0.189	0.209	0.047***	3.997	3.997	0.189	0.189	0.000
<i>Secured</i>	0.507	0.496	0.501	0.5	0.011	0.507	0.507	0.501	0.5	0.000
<i>Guarantor</i>	0.102	0.086	0.303	0.28	0.016	0.102	0.102	0.303	0.303	0.000

APPENDIX B (continued)
Additional Analyses

Panel F: Institutional Lenders' Demand for Private Information and Borrowing Terms - Treatment Sample

	<i>Interest Spread</i>		<i>Amounts</i>		<i>Maturity</i>	
	(1)	(2)	(3)	(4)	(5)	(6)
<i>Post No Inst. Lender</i>	-0.166 (-1.40)	0.018 (0.15)	0.872*** (4.50)	1.012*** (3.40)	0.136** (2.13)	0.148 (1.53)
<i>Had Inst. Lender</i>	-0.428*** (-2.70)	-0.118 (-0.88)	0.698*** (2.90)	0.983*** (2.93)	0.187** (2.13)	0.150 (1.42)
<i>Post No Inst. Lender x Had Inst. Lender</i>	0.385** (2.46)	0.078 (0.55)	-0.925*** (-2.80)	-1.188*** (-2.96)	-0.208** (-2.01)	-0.217* (-1.87)
<i>Assets</i>	-0.008 (-0.20)	0.010 (0.26)	0.231* (1.85)	0.323** (2.58)	0.004 (0.14)	0.025 (1.14)
<i>Current Ratio</i>	0.014 (0.33)	0.051 (1.22)	0.080 (0.82)	0.020 (0.19)	-0.002 (-0.06)	0.005 (0.22)
<i>Leverage</i>	0.118 (0.70)	0.305** (2.52)	-0.036 (-0.06)	0.290 (0.53)	-0.169 (-1.53)	-0.003 (-0.04)
<i>MTB</i>	-0.002 (-1.50)	-0.002 (-1.11)	0.005 (1.34)	0.005 (1.24)	-0.000 (-0.09)	-0.000 (-0.44)
<i>Sales Growth</i>	-0.240* (-1.68)	-0.256** (-2.14)	-0.831** (-2.01)	-0.814** (-2.02)	-0.037 (-0.34)	-0.217** (-2.24)
<i>Interest Coverage</i>	0.000 (1.00)	0.000 (0.35)	-0.000 (-0.77)	-0.000 (-0.39)	0.000 (0.50)	0.000 (0.77)
<i>ROA</i>	-0.263 (-0.82)	0.069 (0.24)	2.023** (2.19)	3.071*** (3.84)	-0.084 (-0.40)	0.306 (1.65)
<i>Altman Z</i>	-0.015 (-1.13)	-0.026** (-2.11)	-0.051 (-1.52)	-0.073** (-2.41)	-0.008 (-0.66)	-0.016 (-1.47)
<i>Age</i>	0.002 (0.46)	0.001 (0.40)	-0.016** (-2.32)	-0.020*** (-2.80)	-0.001 (-0.94)	-0.001 (-0.76)
<i>Past Return</i>	0.162*** (3.07)	0.103* (1.82)	-0.258 (-1.48)	-0.371* (-1.76)	-0.024 (-0.61)	-0.021 (-0.53)
<i>Amounts</i>	-0.007 (-0.22)	-0.041* (-1.76)			0.044** (2.36)	-0.000 (-0.04)
<i>Maturity</i>	-0.023 (-0.31)	-0.011 (-0.10)	0.641** (2.52)	-0.010 (-0.04)		
<i>Secured</i>	0.141*** (2.81)	0.062 (1.30)	0.080 (0.41)	0.121 (0.64)	0.061 (1.49)	0.018 (0.47)
<i>Guarantor</i>	-0.162** (-2.57)	-0.063 (-1.56)	0.198 (1.34)	0.351 (1.62)	0.001 (0.03)	0.051 (1.56)
<i>Model</i>	OLS	OLS	OLS	OLS	OLS	OLS
<i>Firm FE</i>	Yes	Yes	Yes	Yes	Yes	Yes
<i>Year FE</i>	No	No	No	No	No	No
<i>Year-Quarter FE</i>	Yes	Yes	Yes	Yes	Yes	Yes
<i>Credit Rating FE</i>	No	Yes	No	Yes	No	Yes
<i>Loan Type FE</i>	No	Yes	No	Yes	No	Yes
<i>Observations</i>	412	412	412	412	412	412
<i>Adj. (Pseudo) R-squared</i>	0.614	0.755	0.625	0.670	0.235	0.555

APPENDIX B (continued)
Additional Analyses

Panel G: Institutional Lender Migration and Borrower Risk

	<i>Altman Z</i>		<i>Credit Rating</i>		<i>Interest Coverage</i>	
	(1)	(2)	(3)	(4)	(5)	(6)
<i>Treatment Firm x Post No Inv Asset Mgr</i>	-0.241 (-0.81)	-0.346 (-1.12)	-0.178 (-0.22)	-0.201 (-0.25)	-55.126 (-1.01)	-51.817 (-0.93)
<i>Treatment Firm x Post No Inv Asset Mgr x Had Inv Asset Mgr</i>	-0.310 (-0.31)	-0.197 (-0.20)	2.746 (0.79)	2.788 (0.78)	86.745 (0.95)	98.847 (0.88)
<i>Treatment Firm x Post Inv Asset Mgr</i>	0.178 (0.27)	0.198 (0.29)	1.697 (0.64)	1.556 (0.57)	6.356 (0.11)	12.222 (0.17)
<i>Treatment Firm x Post Inv Asset Mgr x Had Inv Asset Mgr</i>	-0.288 (-0.37)	-0.222 (-0.28)	-3.143 (-0.96)	-3.182 (-0.94)	7.347 (0.11)	-23.172 (-0.27)
<i>Post No Inv Asset Mgr</i>	0.491 (1.58)	0.531* (1.72)	1.427 (1.48)	1.496 (1.55)	50.404 (0.78)	50.484 (0.81)
<i>Had Inv Asset Mgr</i>	0.286 (0.94)	0.223 (0.72)	1.811* (1.71)	2.019* (1.95)	55.708 (0.90)	55.557 (0.93)
<i>Post No Inv Asset Mgr x Had Inv Assets</i>	-0.214 (-0.38)	-0.279 (-0.52)	-1.685 (-1.26)	-1.912 (-1.48)	-66.103 (-0.95)	-58.233 (-0.85)
<i>Current Ratio</i>	-1.020*** (-5.73)	-1.022*** (-5.54)	1.398*** (3.28)	1.419*** (3.28)	4.557 (0.38)	15.951 (1.21)
<i>Leverage</i>	0.565*** (3.29)	0.607*** (3.49)	0.617* (1.95)	0.592* (1.91)	32.519** (2.04)	32.647** (2.04)
<i>MTB</i>	-2.545*** (-5.35)	-2.555*** (-5.68)	5.592*** (3.50)	5.625*** (3.54)	-102.984** (-2.06)	-68.704 (-1.54)
<i>Sales Growth</i>	0.015** (2.07)	0.014** (1.97)	-0.008 (-0.67)	-0.009 (-0.73)	-0.247 (-0.49)	-0.280 (-0.53)
<i>Interest Coverage</i>	0.451** (2.24)	0.499** (2.44)	-0.805 (-1.24)	-0.939 (-1.46)	-26.288 (-1.36)	-34.002 (-1.59)
<i>ROA</i>	0.002*** (3.58)	0.002*** (3.68)	-0.002*** (-2.91)	-0.002*** (-2.91)		
<i>Altman Z</i>	1.412* (1.74)	1.717** (2.47)	0.146 (0.08)	-0.036 (-0.02)	8.277 (0.14)	7.645 (0.12)
<i>Age</i>			-0.053 (-0.57)	-0.041 (-0.43)	40.240*** (4.06)	40.450*** (4.12)
<i>Past Return</i>	-0.002 (-0.12)	-0.005 (-0.38)	-0.041 (-1.55)	-0.039 (-1.45)	-1.867* (-1.67)	-1.900* (-1.76)
<i>Amounts</i>	-0.106 (-0.54)	-0.062 (-0.32)	-0.702 (-1.56)	-0.781* (-1.72)	-40.136* (-1.75)	-45.799** (-2.07)
<i>Maturity</i>	0.088 (1.56)	0.111* (1.89)	-0.080 (-0.53)	-0.143 (-0.95)	-10.897* (-1.83)	-11.843* (-1.69)
	-0.187 (-0.77)	-0.154 (-0.53)	0.010 (0.02)	-0.209 (-0.40)	-16.898 (-0.58)	-16.233 (-0.53)

APPENDIX B (continued)
Additional Analyses

Panel G: Institutional Lender Migration and Borrower Risk

	<i>Altman Z</i>		<i>Credit Rating</i>		<i>Interest Coverage</i>	
	(1)	(2)	(3)	(4)	(5)	(6)
<i>Secured</i>	-0.275*	-0.312*	-0.009	-0.018	26.199	23.760
	(-1.73)	(-1.92)	(-0.02)	(-0.05)	(1.16)	(1.11)
<i>Guarantor</i>	-0.056	-0.014	0.498	0.356	-21.032*	-15.949
	(-0.31)	(-0.08)	(0.85)	(0.66)	(-1.75)	(-1.27)
<i>Model</i>	OLS	OLS	OLS	OLS	OLS	OLS
<i>Firm FE</i>	Yes	Yes	Yes	Yes	Yes	Yes
<i>Year FE</i>	No	No	No	No	No	No
<i>Year-Quarter FE</i>	Yes	Yes	Yes	Yes	Yes	Yes
<i>Credit Rating FE</i>	No	Yes	No	No	No	Yes
<i>Loan Type FE</i>	No	Yes	No	Yes	No	Yes
<i>Observations</i>	2,129	2,129	2,129	2,129	2,129	2,129
<i>Adj. (Pseudo) R-squared</i>	0.793	0.797	0.675	0.675	0.667	0.671

FIGURE 1

Parallel Trend of Institutional Lending

This figure plots OLS regression coefficient estimates and two-tailed 90th-percentile confidence intervals based on standard errors clustered at the firm level. I replace *Treatment Firm* \times *Post* in Model (1) with separate interactions between *Treatment Firm* and trend variables, each of which equals 1 for every two-year sample period before and after the initiation of the satellite data coverage (and 0 otherwise). The last two-year period before the release of the satellite data (from the fourth quarter of 2013 to the third quarter of 2015) serves as a benchmark.

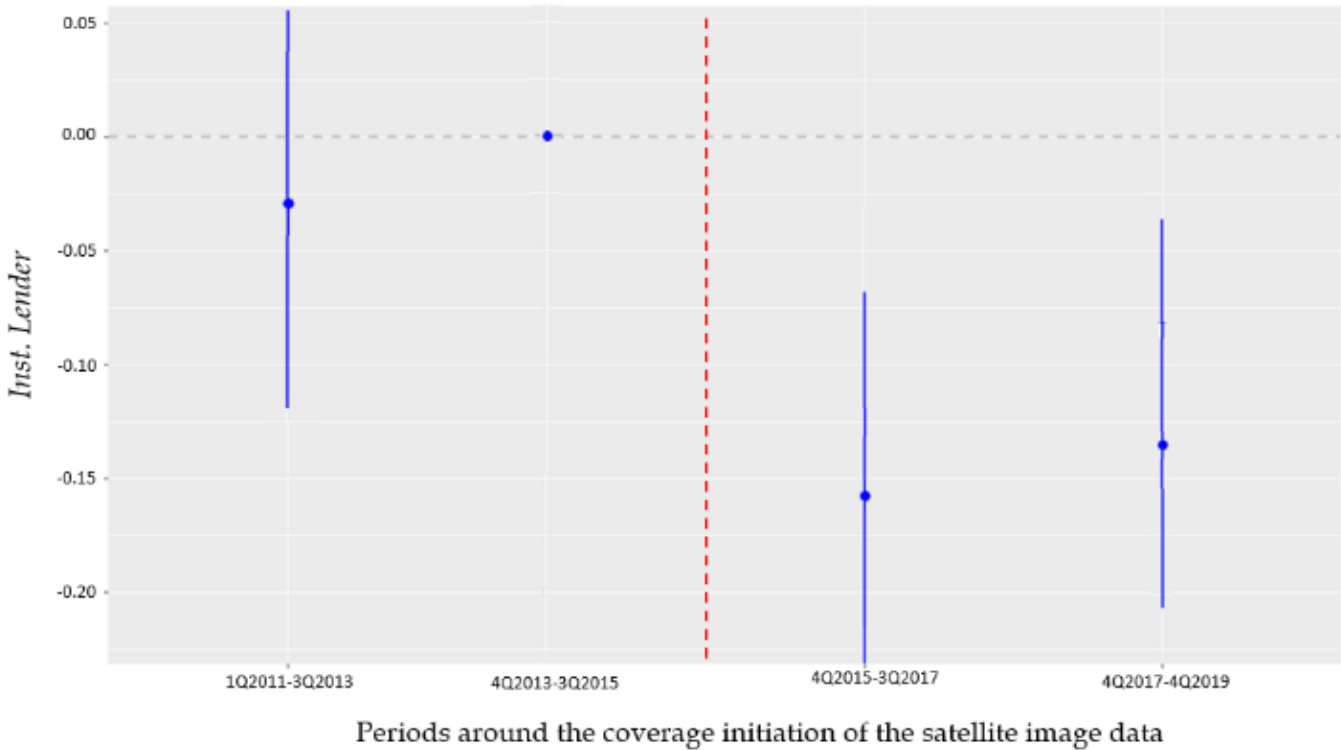


TABLE 1
Descriptive Statistics

This table provides descriptive statistics for the main variables of interest. All variables are defined in Appendix A.

	N	Mean	Median	SD
<i>Inst. Lender</i>	2,129	0.165	0.000	0.372
<i>Post</i>	2,129	0.332	0.000	0.471
<i>Treatment Firm</i>	2,129	0.194	0.000	0.395
<i>Post Inst. Lender</i>	2,129	0.115	0.000	0.320
<i>Post No Inst. Lender</i>	2,129	0.217	0.000	0.412
<i>Had Inst. Lender</i>	2,129	0.104	0.000	0.305
<i>Assets</i>	2,129	8.064	7.961	1.787
<i>Current Ratio</i>	2,129	1.598	1.342	0.978
<i>Leverage</i>	2,129	0.348	0.319	0.267
<i>MTB</i>	2,129	3.230	2.506	9.043
<i>Sales Growth</i>	2,129	0.148	0.070	0.348
<i>Interest Coverage</i>	2,129	65.500	8.409	286.798
<i>ROA</i>	2,129	0.027	0.042	0.120
<i>Altman Z</i>	2,129	3.546	3.290	2.851
<i>Age</i>	2,129	24.178	21.000	18.629
<i>Past Return</i>	2,129	0.055	0.040	0.252
<i>No Analyst Coverage</i>	2,129	0.508	1.000	0.500
<i>No Earnings Forecast</i>	2,129	0.307	0.000	0.461
<i>High Reputation</i>	2,129	0.281	0.000	0.450
<i>Amounts</i>	2,129	19.821	19.808	1.045
<i>Maturity</i>	2,129	3.959	4.096	0.206
<i>Secured</i>	2,129	0.498	0.000	0.500
<i>Guarantor</i>	2,129	0.089	0.000	0.284
<i>Interest Spread</i>	2,129	5.273	5.267	0.339

TABLE 2
Satellite Image Data and Institutional Lending

This table examines whether the probability that institutional lenders issue a loan is lower after the satellite data on a borrower becomes commercially available. Panel A presents the results of univariate analysis. Panel B shows the results of multivariate analysis. Panel C reports the results of analysis using an entropy balancing approach. In Panels B and C, Column(s) 1 (2 and 3) present(s) the results using a logit (OLS) model. t-statistics in parentheses are based on standard errors clustered at the firm level. ***, **, and * indicate significance at the 0.01, 0.05, and 0.10 level, respectively. All variables are defined in Appendix A.

Panel A: Univariate Analysis

<i>Inst. Lender</i>	<i>Treatment Firm=0</i> (a)	<i>Treatment Firm=1</i> (b)	<i>Difference</i> (b) - (a)
<i>Post = 0</i> (c)	0.163	0.160	-0.003
<i>Post = 1</i> (d)	0.194	0.087	-0.107***
<i>Difference</i> (d) - (c)	0.031	-0.073**	-0.104**

Panel B: Multivariate Analysis

	<i>Inst. Lender</i>		
	(1)	(2)	(3)
<i>Treatment Firm x Post</i>	-1.123** (-2.43)	-0.107*** (-2.82)	-0.138*** (-3.48)
<i>Assets</i>	-0.311 (-1.35)	-0.021 (-1.09)	-0.006 (-0.34)
<i>Current Ratio</i>	0.118 (0.58)	0.011 (0.71)	0.012 (0.77)
<i>Leverage</i>	-0.585 (-0.98)	-0.043 (-0.61)	-0.051 (-0.67)
<i>MTB</i>	0.008 (0.79)	0.001 (1.00)	0.001 (1.00)
<i>Sales Growth</i>	0.332 (0.99)	0.031 (0.74)	0.033 (0.79)
<i>Interest Coverage</i>	-0.001* (-1.82)	-0.000 (-1.56)	-0.000* (-1.72)
<i>ROA</i>	-1.790* (-1.69)	-0.160 (-1.41)	-0.180 (-1.52)
<i>Altman Z</i>	-0.083 (-1.41)	-0.010 (-1.60)	-0.012** (-2.00)
<i>Age</i>	-0.001 (-0.03)	-0.000 (-0.00)	-0.001 (-0.44)
<i>Past Return</i>	-0.545 (-1.34)	-0.059 (-1.57)	-0.079** (-2.16)
<i>Amounts</i>	0.206 (1.58)	0.023 (1.60)	0.032** (2.03)

TABLE 2 (continued)
Satellite Image Data and Institutional Lending

Panel B: Multivariate Analysis (continued)

	<i>Inst. Lender</i>		
	(1)	(2)	(3)
<i>Maturity</i>	0.685 (1.32)	0.061 (1.18)	0.092 (1.60)
<i>Secured</i>	0.610** (2.21)	0.075** (2.32)	0.063* (1.85)
<i>Guarantor</i>	0.607 (1.39)	0.052 (1.00)	0.070 (1.27)
<i>Model</i>	Logit	OLS	OLS
<i>Firm FE</i>	Yes	Yes	Yes
<i>Year FE</i>	Yes	Yes	No
<i>Year-Quarter FE</i>	No	No	Yes
<i>Credit Rating FE</i>	No	No	Yes
<i>Loan Type FE</i>	No	No	Yes
<i>Observations</i>	904	2,129	2,129
<i>Adj. (Pseudo) R-squared</i>	0.069	0.230	0.247

Panel C: Entropy Balancing Approach

	<i>Inst. Lender</i>		
	(1)	(2)	(3)
<i>Treatment Firm x Post</i>	-1.787*** (-2.94)	-0.113*** (-2.82)	-0.125*** (-3.16)
<i>Model</i>	Logit	OLS	OLS
<i>Controls</i>	Yes	Yes	Yes
<i>Firm FE</i>	Yes	Yes	Yes
<i>Year FE</i>	Yes	Yes	No
<i>Year-Quarter FE</i>	No	No	Yes
<i>Credit Rating FE</i>	No	No	Yes
<i>Loan Type FE</i>	No	No	Yes
<i>Observations</i>	904	2,129	2,129
<i>Adj. (Pseudo) R-squared</i>	N/A	0.483	0.534

TABLE 3
Falsification Test

This table reports the results of falsification tests using different types of institutional lender as a dependent variable. Panel A presents the results of the falsification test based on OLS and logit model. Columns 1 and 3 (2 and 4) present the results using a logit (OLS) model. In Columns 1 and 2, the dependent variable is *Inst. Lender Bank Affiliated* which equals 1 if the loan is issued with at least one bank-affiliated institutional lender but is not issued with an independent institutional lender, and 0 otherwise. In Columns 3 and 4, the dependent variable is *Inst. Lender No Investment*, which equals 1 if the loan is issued with at least one independent institutional lender that does *not* engage in investment businesses but is not issued with an independent institutional lender that does engage in investment businesses, and 0 otherwise. Panel B reports the results of falsification tests using a multinomial logit model. t-statistics in parentheses are based on standard errors clustered at the firm level. ***, **, and * indicate significance at the 0.01, 0.05, and 0.10 level, respectively. All variables are defined in Appendix A.

Panel A: Falsification Test

	<i>Inst. Lender Bank Affiliated</i>		<i>Inst. Lender No Investment</i>	
	(1)	(2)	(3)	(4)
<i>Treatment Firm x Post</i>	-0.317	-0.024	1.861	0.014
	(-0.68)	(-0.38)	(1.62)	(0.48)
<i>Assets</i>	-0.160	-0.035	0.220	0.021
	(-0.69)	(-1.21)	(0.52)	(1.02)
<i>Current Ratio</i>	-0.237	-0.016	-1.480***	-0.013*
	(-1.23)	(-0.87)	(-4.53)	(-1.78)
<i>Leverage</i>	0.031	-0.010	-1.239	-0.031
	(0.05)	(-0.13)	(-0.84)	(-0.89)
<i>MTB</i>	-0.005	-0.001	0.004	0.000
	(-0.51)	(-0.52)	(0.20)	(0.11)
<i>Sales Growth</i>	-0.204	-0.031	-1.033	-0.028*
	(-0.59)	(-0.74)	(-1.01)	(-1.74)
<i>Interest Coverage</i>	-0.000	-0.000	0.001	-0.000
	(-0.56)	(-0.50)	(0.68)	(-0.13)
<i>ROA</i>	0.245	0.024	-3.575	-0.104
	(0.25)	(0.21)	(-1.28)	(-1.62)
<i>Altman Z</i>	-0.048	-0.009	0.305	0.003
	(-0.74)	(-1.03)	(1.53)	(1.05)
<i>Age</i>	-0.014	-0.002	0.012	0.002*
	(-0.96)	(-0.90)	(0.49)	(1.82)
<i>Past Return</i>	0.324	0.050	-0.809	-0.027
	(1.02)	(1.19)	(-1.40)	(-1.17)
<i>Amounts</i>	0.285**	0.030*	0.037	0.013
	(2.32)	(1.70)	(0.13)	(1.35)

TABLE 3 (continued)
Falsification Test

Panel A: Falsification Test (continued)

	<i>Inst. Lender Bank Affiliated</i>		<i>Inst. Lender No Investment</i>	
	(1)	(2)	(3)	(4)
<i>Maturity</i>	0.642 (1.62)	0.077 (1.23)	-0.766 (-0.62)	-0.044 (-1.11)
<i>Secured</i>	-0.216 (-0.78)	-0.023 (-0.63)	0.966 (1.29)	0.050*** (2.67)
<i>Guarantor</i>	0.155 (0.43)	0.028 (0.60)	-0.109 (-0.19)	-0.010 (-0.31)
<i>Model</i>	Logit	OLS	Logit	OLS
<i>Firm FE</i>	Yes	Yes	Yes	Yes
<i>Year FE</i>	Yes	No	Yes	No
<i>Year-Quarter FE</i>	No	Yes	No	Yes
<i>Credit Rating FE</i>	No	Yes	No	Yes
<i>Loan Type FE</i>	No	Yes	No	Yes
<i>Observations</i>	1,043	2,129	295	2,129
<i>Adj. (Pseudo) R-squared</i>	0.029	0.248	0.291	0.349

Panel B: Falsification Test—Multinomial Logit Model Analysis

	<i>Inst. Lender Bank Affiliated = 1</i>	<i>Inst. Lender No Investment = 1</i>	<i>Inst. Lender = 1</i>
	(1)	(2)	(3)
<i>Treatment Firm x Post</i>	-0.894 (-1.36)	0.807 (0.67)	-2.197*** (-2.86)
<i>Assets</i>	-0.226 (-0.62)	0.312 (0.46)	-0.302 (-0.86)
<i>Current Ratio</i>	-0.236 (-0.87)	-0.604 (-0.68)	0.008 (0.03)
<i>Leverage</i>	-0.152 (-0.18)	-0.781 (-0.42)	-0.622 (-0.66)
<i>MTB</i>	0.001 (0.10)	0.021 (0.58)	0.012 (0.74)
<i>Sales Growth</i>	-0.263 (-0.53)	-1.593 (-1.02)	0.245 (0.54)
<i>Interest Coverage</i>	-0.001 (-1.20)	0.000 (0.22)	-0.002** (-1.99)
<i>ROA</i>	-0.859 (-0.53)	-5.589 (-1.50)	-2.767 (-1.50)
<i>Altman Z</i>	-0.095 (-1.07)	0.116 (0.43)	-0.100 (-1.20)

TABLE 3 (continued)
Falsification Test

Panel B: Falsification Test—Multinomial Logit Model Analysis (continued)

	<i>Inst. Lender Bank Affiliated = 1</i>	<i>Inst. Lender No Investment = 1</i>	<i>Inst. Lender = 1</i>
	(1)	(2)	(3)
<i>AGE</i>	-0.010 (-0.34)	0.005 (0.10)	-0.004 (-0.10)
<i>Past Return</i>	-0.026 (-0.05)	-0.694 (-0.86)	-0.822 (-1.37)
<i>Amounts</i>	0.345** (1.96)	0.264 (0.67)	0.397** (2.11)
<i>Maturity</i>	0.711 (1.34)	-0.262 (-0.16)	0.799 (1.10)
<i>Secured</i>	0.115 (0.28)	1.005 (1.05)	0.758* (1.81)
<i>Guarantor</i>	0.558 (0.98)	0.135 (0.16)	1.086 (1.56)
<i>Model</i>	Mlogit	Mlogit	Mlogit
<i>Controls</i>	Yes	Yes	Yes
<i>Firm FE</i>	Yes	Yes	Yes
<i>Year FE</i>	Yes	Yes	Yes
<i>Year-Quarter FE</i>	No	No	No
<i>Credit Rating FE</i>	No	No	No
<i>Loan Type FE</i>	No	No	No
<i>Observations</i>	2,129	2,129	2,129
<i>Adj. (Pseudo) R-squared</i>	0.617	0.617	0.617
Test: [1] <i>Treatment Firm x Post</i> > [3] <i>Treatment Firm x Post</i>	p-value: 0.040		
Test: [2] <i>Treatment Firm x Post</i> > [3] <i>Treatment Firm x Post</i>	p-value: 0.011		

TABLE 4
Borrower Opacity

This table examines whether the effect of the satellite image data coverage on institutional lender participation is more pronounced when borrowers are opaque. Panels A, B and C report the results of the analyses in which borrower opacity is measured by, respectively, a borrower's equity analyst coverage (*No Analyst Coverage*), whether a borrower issues earnings forecasts (*No Earnings Forecast*), and a borrower's press releases (*Low Press Releases*). In all panels, Columns 1 and 2 (3 and 4) present the results using a logit (OLS) model. t-statistics in parentheses are based on standard errors clustered at the firm level. ***, **, and * indicate significance at the 0.01, 0.05, and 0.10 level, respectively. All variables are defined in Appendix A.

Panel A: No Analyst Coverage

	<i>Inst. Lender</i>			
	<i>No Analyst Coverage=0</i>	<i>No Analyst Coverage=1</i>	<i>No Analyst Coverage=0</i>	<i>No Analyst Coverage=1</i>
	(1)	(2)	(3)	(4)
<i>Treatment Firm x Post</i>	-0.426 (-0.79)	-2.613*** (-2.91)	-0.071 (-1.42)	-0.254*** (-3.79)
<i>Model</i>	Logit	Logit	OLS	OLS
<i>Controls</i>	Yes	Yes	Yes	Yes
<i>Firm FE</i>	Yes	Yes	Yes	Yes
<i>Year FE</i>	Yes	Yes	No	No
<i>Year-Quarter FE</i>	No	No	Yes	Yes
<i>Credit Rating FE</i>	No	No	Yes	Yes
<i>Loan Type FE</i>	No	No	Yes	Yes
<i>Observations</i>	492	412	1,047	1,082
<i>Adj. (Pseudo) R-squared</i>	0.105	0.140	0.252	0.251
Test: [1] <i>Treatment Firm x Post</i> > [2] <i>Treatment Firm x Post</i>	p-value: 0.012			
Test: [3] <i>Treatment Firm x Post</i> > [4] <i>Treatment Firm x Post</i>	p-value: 0.007			

TABLE 4 (continued)
Borrower Opacity

Panel B: No Earnings Forecast

	<i>Inst. Lender</i>			
	<i>No Earnings Forecast=0</i>	<i>No Earnings Forecast=1</i>	<i>No Earnings Forecast=0</i>	<i>No Earnings Forecast=1</i>
	(1)	(2)	(3)	(4)
<i>Treatment Firm x Post</i>	-0.809* (-1.75)	-16.590*** (-13.53)	-0.113*** (-2.77)	-0.349** (-2.40)
<i>Model</i>	Logit	Logit	OLS	OLS
<i>Controls</i>	Yes	Yes	Yes	Yes
<i>Firm FE</i>	Yes	Yes	Yes	Yes
<i>Year FE</i>	Yes	Yes	No	No
<i>Year-Quarter FE</i>	No	No	Yes	Yes
<i>Credit Rating FE</i>	No	No	Yes	Yes
<i>Loan Type FE</i>	No	No	Yes	Yes
<i>Observations</i>	649	255	1,475	654
<i>Adj. (Pseudo) R-squared</i>	0.066	0.246	0.255	0.222
Test: [1] <i>Treatment Firm x Post</i> > [2] <i>Treatment Firm x Post</i>	p-value: 0.042			
Test: [3] <i>Treatment Firm x Post</i> > [4] <i>Treatment Firm x Post</i>	p-value: 0.030			

Panel C: Low Press Releases

	<i>Inst. Lender</i>			
	<i>Low Press Releases=0</i>	<i>Low Press Releases=1</i>	<i>Low Press Releases=0</i>	<i>Low Press Releases=1</i>
	(1)	(2)	(3)	(4)
<i>Treatment Firm x Post</i>	-0.824 (-1.54)	-1.997** (-2.33)	-0.091 (-1.55)	-0.187*** (-3.33)
<i>Model</i>	Logit	Logit	OLS	OLS
<i>Controls</i>	Yes	Yes	Yes	Yes
<i>Firm FE</i>	Yes	Yes	Yes	Yes
<i>Year FE</i>	Yes	Yes	No	No
<i>Year-Quarter FE</i>	No	No	Yes	Yes
<i>Credit Rating FE</i>	No	No	Yes	Yes
<i>Loan Type FE</i>	No	No	Yes	Yes
<i>Observations</i>	551	353	1,065	1,064
<i>Adj. (Pseudo) R-squared</i>	0.080	0.149	0.197	0.319
Test: [1] <i>Treatment Firm x Post</i> > [2] <i>Treatment Firm x Post</i>	p-value: 0.097			
Test: [3] <i>Treatment Firm x Post</i> > [4] <i>Treatment Firm x Post</i>	p-value: 0.099			

TABLE 5

Early Dissemination of Borrower Private Information

This table examines whether early dissemination of borrower private information is important to the relationship between the satellite image data coverage and institutional lender participation. Panels A and B report the results of the analyses in which the borrower’s information dissemination is measured by the number of performance covenants in the loan (*High Perf. Covenants*) and by the lender’s reputation (*High Reputation*), respectively. In all panels, Columns 1 and 2 (3 and 4) present the results using a logit (OLS) model. t-statistics in parentheses are based on standard errors clustered at the firm level. ***, **, and * indicate significance at the 0.01, 0.05, and 0.10 level, respectively. All variables are defined in Appendix A.

Panel A: High Performance Covenants

	<i>Inst. Lender</i>			
	<i>High Perf. Covenants=0</i>	<i>High Perf. Covenants=1</i>	<i>High Perf. Covenants=0</i>	<i>High Perf. Covenants=1</i>
	(1)	(2)	(3)	(4)
<i>Treatment Firm x Post</i>	-0.301 (-0.35)	-1.828*** (-3.01)	-0.088 (-1.24)	-0.187*** (-3.82)
<i>Model</i>	Logit	Logit	OLS	OLS
<i>Controls</i>	Yes	Yes	Yes	Yes
<i>Firm FE</i>	Yes	Yes	Yes	Yes
<i>Year FE</i>	Yes	Yes	No	No
<i>Year-Quarter FE</i>	No	No	Yes	Yes
<i>Credit Rating FE</i>	No	No	Yes	Yes
<i>Loan Type FE</i>	No	No	Yes	Yes
<i>Observations</i>	381	523	1,030	1,099
<i>Adj. (Pseudo) R-squared</i>	0.090	0.152	0.242	0.256
Test: [1] <i>Treatment Firm x Post</i> > [2] <i>Treatment Firm x Post</i>	p-value: 0.051			
Test: [3] <i>Treatment Firm x Post</i> > [4] <i>Treatment Firm x Post</i>	p-value: 0.094			

TABLE 5 (continued)
Early Dissemination of Borrower Private Information

Panel B: Lender Reputation

	<i>Inst. Lender</i>			
	<i>High Reputation=0</i>	<i>High Reputation=1</i>	<i>High Reputation=0</i>	<i>High Reputation=1</i>
	(1)	(2)	(3)	(4)
<i>Treatment Firm x Post</i>	-0.663 (-1.00)	-2.078*** (-3.51)	-0.078 (-1.59)	-0.294*** (-3.76)
<i>Model</i>	Logit	Logit	OLS	OLS
<i>Controls</i>	Yes	Yes	Yes	Yes
<i>Firm FE</i>	Yes	Yes	Yes	Yes
<i>Year FE</i>	Yes	Yes	No	No
<i>Year-Quarter FE</i>	No	No	Yes	Yes
<i>Credit Rating FE</i>	No	No	Yes	Yes
<i>Loan Type FE</i>	No	No	Yes	Yes
<i>Observations</i>	586	318	1,531	598
<i>Adj. (Pseudo) R-squared</i>	0.061	0.173	0.258	0.223
Test: [1] <i>Treatment Firm x Post</i> > [2] <i>Treatment Firm x Post</i>	p-value: 0.057			
Test: [3] <i>Treatment Firm x Post</i> > [4] <i>Treatment Firm x Post</i>	p-value: 0.005			

TABLE 6
Accuracy of Satellite Image Data

This table examines whether the effect of the satellite image data coverage on institutional lending is greater when the data predicts borrowers' performance more accurately. Panels A and B report the results of analyses in which the accuracy of the satellite image data is measured by the correlation between quarterly changes in store car counts and quarterly changes in the borrower's sales (*Treatment Firm High Corr*, *Treatment Firm Low Corr*), and the average standard deviation of quarterly changes in car counts across stores (*Treatment Firm High SD*, *Treatment Firm Low SD*), respectively. In all panels, Column(s) 1 (3 and 4) present(s) the results using a logit (OLS) model. t-statistics in parentheses are based on standard errors clustered at the firm level. ***, **, and * indicate significance at the 0.01, 0.05, and 0.10 level, respectively. All variables are defined in Appendix A.

Panel A: Accuracy of the Satellite Image Data—High Correlations

	<i>Inst. Lender</i>		
	(1)	(2)	(3)
<i>Treatment Firm High Corr x Post</i>	-1.852*** (-2.78)	-0.176*** (-3.82)	-0.194*** (-4.16)
<i>Treatment Firm Low Corr x Post</i>	-0.451 (-0.80)	-0.067 (-1.28)	-0.080 (-1.54)
<i>Model</i>	Logit	OLS	OLS
<i>Controls</i>	Yes	Yes	Yes
<i>Firm FE</i>	Yes	Yes	Yes
<i>Year FE</i>	Yes	Yes	No
<i>Year-Quarter FE</i>	No	No	Yes
<i>Credit Rating FE</i>	No	No	Yes
<i>Loan Type FE</i>	No	No	Yes
<i>Observations</i>	904	2,129	2,129
<i>Adj. (Pseudo) R-squared</i>	0.073	0.239	0.248
Test: <i>Treatment Firm Low Corr x Post</i> > <i>Treatment Firm High Corr x Post</i>	p-value: 0.042	p-value: 0.038	p-value: 0.029

Panel B: Accuracy of the Satellite Image Data—Low Standard Deviation

	<i>Inst. Lender</i>		
	(1)	(2)	(3)
<i>Treatment Firm High SD x Post</i>	-0.654 (-1.13)	-0.082* (-1.66)	-0.094* (-1.92)
<i>Treatment Firm Low SD x Post</i>	-1.572*** (-2.61)	-0.165*** (-3.34)	-0.184*** (-3.59)
<i>Model</i>	Logit	OLS	OLS
<i>Controls</i>	Yes	Yes	Yes
<i>Firm FE</i>	Yes	Yes	Yes
<i>Year FE</i>	Yes	Yes	No
<i>Year-Quarter FE</i>	No	No	Yes
<i>Credit Rating FE</i>	No	No	Yes
<i>Loan Type FE</i>	No	No	Yes
<i>Observations</i>	904	2,129	2,129
<i>Adj. (Pseudo) R-squared</i>	0.071	0.238	0.247
Test: <i>Treatment Firm High SD x Post</i> > <i>Treatment Firm Low SD x Post</i>	p-value: 0.115	p-value: 0.094	p-value: 0.075

TABLE 7
Institutional Lenders' Demand for Private Information and Borrowing Terms

This table examines whether institutional lenders' information demand affects borrowers' credit outcomes. The dependent variable is *Interest Spread* in Columns 1 and 2, *Amounts* in Columns 3 and 4, and *Maturity* Columns 5 and 6. The main variable of interest is *Treatment Firm x Post No Inst. Lender x Had Inst. Lender*, which captures loans issued to treatment borrowers (*Treatment Firm* = 1) who do not obtain loans from institutional lenders in the coverage period (*Post No Inst. Lender* = 1) but had lending relationships with institutional lenders in the pre-coverage period (*Had Inst. Lender* = 1). t-statistics in parentheses are based on standard errors clustered at the firm level. ***, **, and * indicate significance at the 0.01, 0.05, and 0.10 level, respectively. All variables are defined in Appendix A.

	<i>Interest Spread</i>		<i>Amounts</i>		<i>Maturity</i>	
	(1)	(2)	(3)	(4)	(5)	(6)
<i>Treatment Firm x Post No Inst. Lender</i>	0.038 (1.04)	0.029 (0.77)	0.156 (1.45)	0.146 (1.32)	0.001 (0.04)	0.015 (0.52)
<i>Treatment Firm x Post No Inst. Lender x Had Inst. Lender</i>	0.348*** (2.68)	0.233* (1.72)	-0.580* (-1.72)	-0.596* (-1.72)	-0.237** (-2.35)	-0.174* (-1.90)
<i>Treatment Firm x Post Inst. Lender</i>	0.089 (0.99)	0.048 (0.50)	-0.716*** (-3.80)	-0.704*** (-3.42)	-0.102* (-1.73)	-0.092 (-1.39)
<i>Treatment Firm x Post Inst. Lender x Had Inst. Lender</i>	-0.393*** (-3.30)	-0.284** (-2.38)	0.605** (2.43)	0.639** (2.49)	0.243*** (2.97)	0.157* (1.91)
<i>Post No Inst. Lender</i>	-0.075 (-1.43)	-0.037 (-0.77)	-0.100 (-0.93)	-0.132 (-1.33)	0.039 (1.19)	0.027 (0.98)
<i>Had Inst. Lender</i>	-0.009 (-0.17)	0.016 (0.35)	-0.057 (-0.34)	-0.093 (-0.73)	-0.028 (-0.77)	-0.002 (-0.06)
<i>Post No Inst. Lender x Had Inst. Lender</i>	-0.000 (-0.00)	-0.022 (-0.35)	-0.158 (-0.76)	-0.077 (-0.46)	0.040 (0.82)	0.000 (0.01)
<i>Assets</i>	-0.038* (-1.92)	-0.020 (-1.10)	0.273*** (5.32)	0.246*** (4.78)	-0.017 (-1.33)	-0.018 (-1.35)
<i>Current Ratio</i>	-0.002 (-0.19)	0.004 (0.37)	-0.031 (-1.00)	-0.037 (-1.31)	0.013 (1.20)	0.005 (0.53)
<i>Leverage</i>	0.035 (0.66)	0.055 (1.00)	0.075 (0.47)	0.128 (0.93)	-0.065 (-1.54)	-0.077* (-1.86)
<i>MTB</i>	-0.001 (-1.44)	-0.001 (-1.50)	0.003 (1.41)	0.002 (1.08)	0.000 (0.30)	0.000 (0.98)
<i>Sales Growth</i>	-0.063** (-2.36)	-0.044* (-1.83)	0.025 (0.34)	-0.066 (-1.02)	-0.020 (-0.98)	-0.029 (-1.48)
<i>Interest Coverage</i>	-0.000 (-0.52)	-0.000 (-0.54)	-0.000** (-2.02)	-0.000* (-1.81)	-0.000 (-0.59)	-0.000 (-0.54)
<i>ROA</i>	-0.160* (-1.78)	-0.142 (-1.64)	0.828*** (3.36)	0.832*** (3.52)	0.073 (0.96)	0.062 (0.95)

TABLE 7 (continued)
Institutional Lenders' Demand for Private Information and Borrowing Terms

	<i>Interest Spread</i>		<i>Amounts</i>		<i>Maturity</i>	
	(1)	(2)	(3)	(4)	(5)	(6)
<i>Altman Z</i>	-0.005 (-0.75)	-0.007 (-1.17)	0.019* (1.65)	0.023** (2.00)	-0.003 (-0.78)	-0.002 (-0.53)
<i>Age</i>	0.003** (2.02)	0.002 (1.36)	-0.009** (-2.41)	-0.007* (-1.73)	0.001 (0.38)	0.001 (0.45)
<i>Past Return</i>	0.050* (1.88)	0.039 (1.54)	0.013 (0.17)	0.002 (0.03)	0.004 (0.18)	0.009 (0.52)
<i>Amounts</i>	-0.033** (-2.29)	-0.051*** (-4.28)			0.041*** (4.93)	0.035*** (4.44)
<i>Maturity</i>	-0.124*** (-3.26)	-0.036 (-0.89)	0.494*** (5.08)	0.509*** (4.53)		
<i>Secured</i>	0.120*** (5.09)	0.068*** (3.32)	0.064 (1.02)	0.009 (0.15)	0.055*** (3.30)	0.018 (1.10)
<i>Guarantor</i>	-0.118*** (-3.65)	-0.051* (-1.93)	0.259** (1.99)	0.187* (1.67)	0.004 (0.14)	-0.009 (-0.36)
<i>Model</i>	OLS	OLS	OLS	OLS	OLS	OLS
<i>Firm FE</i>	Yes	Yes	Yes	Yes	Yes	Yes
<i>Year FE</i>	No	No	No	No	No	No
<i>Year-Quarter FE</i>	Yes	Yes	Yes	Yes	Yes	Yes
<i>Credit Rating FE</i>	No	Yes	No	Yes	No	Yes
<i>Loan Type FE</i>	No	Yes	No	Yes	No	Yes
<i>Observations</i>	2,129	2,129	2,129	2,129	2,129	2,129
<i>Adj. (Pseudo) R-squared</i>	0.585	0.681	0.658	0.694	0.263	0.454

TABLE 8

Availability of Satellite Image Data and Alternative Sources of Capital

This table examines whether the enhanced transparency facilitated by the availability of the satellite data can reduce costs of raising capital other than loan issuance. Panel A reports the results of analyses that examine whether the initiation of the satellite data coverage reduces a borrower’s cost of equity capital. In Columns 1 to 4, I employ various internal rate of returns that equates a firm’s forecasted cash flows to its current market price: *AGR ICC* following Ohlson and Juettner-Nauroth (2005) in Column 1, *CAT ICC* following Claus and Thomas (2001) in Column 2, *GLS ICC* following Gebhardt et al. (2001) in Column 3, and *PEG ICC* following Easton (2004) in Column 4. In Column 5, *AVG ICC* is an equally-weighted average of the following four measures of cost of capital: *AGR* (Ohlson and Juettner-Nauroth 2005), *CAT* (Claus and Thomas 2001), *GLS* (Gebhardt et al. 2001), and *PEG* (Easton 2004). Panel B employs *AVG ICC* as a cost of capital proxy and investigates whether the satellite data coverage reduces cost of equity capital for opaque borrowers. In Columns 1, 2, and 3, I use *No Analyst Coverage*, *No Earnings Forecast*, and *Low Press Releases* to capture opaque borrowers respectively. Panel C presents the results of analyses that examine whether a borrower is more likely to issue equity in the coverage period. In Columns 1 and 2, *Equity Issuance Amount* is the natural logarithm of equity amounts raised by the borrower, measured in the year of the loan’s issuance. In Columns 3, 4, and 5, *Equity Issuance Indicator* is an indicator variable equal to 1 if the borrower issues equity in the year of the loan’s issuance, and 0 otherwise. Panel D examines whether opaque borrowers are more likely to issue equity in the coverage period. To capture opaque borrowers, I use *No Analyst Coverage* in Columns 1, 4, *No Earnings Forecast* in Columns 2, 5, and *Low Press Releases* in Columns 3, 6. t-statistics in parentheses are based on standard errors clustered at the firm level. ***, **, and * indicate significance at the 0.01, 0.05, and 0.10 level, respectively. All variables are defined in Appendix A.

Panel A: Availability of Satellite Image Data and Cost of Equity Capital

	<i>AGR ICC</i>	<i>CAT ICC</i>	<i>GLS ICC</i>	<i>PEG ICC</i>	<i>AVG ICC</i>
	(1)	(2)	(3)	(4)	(5)
<i>Treatment Firm x Post</i>	0.038 (0.95)	0.009 (1.59)	0.013 (1.04)	0.003 (0.33)	0.018* (1.85)
<i>Assets</i>	0.019 (1.23)	-0.005 (-1.08)	-0.020** (-2.38)	-0.004 (-0.70)	-0.008* (-1.75)
<i>Current Ratio</i>	-0.018 (-1.45)	-0.002 (-0.48)	-0.007 (-1.60)	-0.002 (-0.27)	-0.008** (-2.05)
<i>Leverage</i>	0.012 (0.24)	0.006 (0.44)	-0.053** (-1.96)	0.013 (0.49)	-0.008 (-0.42)
<i>MTB</i>	-0.001 (-0.58)	0.000 (0.54)	-0.001* (-1.69)	-0.000 (-0.87)	-0.001*** (-3.05)
<i>Sales Growth</i>	-0.045 (-1.58)	-0.006 (-0.81)	0.027* (1.84)	-0.031*** (-3.20)	0.001 (0.07)
<i>Interest Coverage</i>	-0.000 (-0.52)	0.000 (0.03)	0.000 (0.08)	0.000 (1.53)	0.000 (0.14)
<i>ROA</i>	-0.477*** (-4.90)	0.172*** (7.16)	-0.067 (-1.45)	-0.430*** (-8.74)	-0.051* (-1.81)

TABLE 8 (continued)
Availability of Satellite Image Data and Alternative Source of Capital

Panel A: Availability of Satellite Image Data and Cost of Equity Capital (continued)

	<i>AGR ICC</i>	<i>CAT ICC</i>	<i>GLS ICC</i>	<i>PEG ICC</i>	<i>AVG ICC</i>
	(1)	(2)	(3)	(4)	(5)
<i>Altman Z</i>	0.010** (2.24)	-0.003*** (-2.80)	-0.004** (-2.57)	0.000 (0.14)	-0.001 (-0.48)
<i>AGE</i>	0.002 (0.95)	-0.001** (-2.08)	0.002** (2.46)	-0.001 (-1.00)	0.000 (1.05)
<i>Past Return</i>	0.002 (0.07)	-0.010 (-1.56)	0.013 (1.11)	-0.013 (-1.42)	0.000 (0.06)
<i>Amounts</i>	-0.002 (-0.21)	0.001 (0.67)	-0.005 (-1.55)	-0.004 (-1.42)	-0.001 (-0.56)
<i>Maturity</i>	0.015 (0.48)	0.007 (1.04)	-0.000 (-0.00)	-0.000 (-0.01)	-0.005 (-0.55)
<i>Secured</i>	-0.003 (-0.21)	-0.006* (-1.67)	0.009 (1.15)	-0.008 (-1.31)	0.004 (1.00)
<i>Guarantor</i>	0.011 (0.35)	-0.010* (-1.79)	0.018 (1.44)	-0.013 (-1.33)	0.006 (0.86)
<i>Model</i>	OLS	OLS	OLS	OLS	OLS
<i>Firm FE</i>	Yes	Yes	Yes	Yes	Yes
<i>Year FE</i>	No	No	No	No	No
<i>Year-Quarter FE</i>	Yes	Yes	Yes	Yes	Yes
<i>Credit Rating FE</i>	Yes	Yes	Yes	Yes	Yes
<i>Loan Type FE</i>	Yes	Yes	Yes	Yes	Yes
<i>Observations</i>	2,129	2,129	2,129	2,129	2,129
<i>Adj. (Pseudo) R-squared</i>	0.717	0.339	0.442	0.687	0.441

Panel B: Availability of Satellite Image Data and Cost of Equity Capital—Opaque Borrowers

	<i>AVG ICC</i>		
	<i>No Analyst Coverage</i>	<i>No Earnings Forecast</i>	<i>Low Press Releases</i>
	(1)	(2)	(3)
<i>Treatment Firm x Post</i>	0.032** (2.33)	0.024** (2.04)	0.026* (1.68)
<i>Post x Opaque Borrowers</i>	-0.003 (-0.31)	0.011 (1.15)	-0.001 (-0.06)
<i>Treatment Firm x Post x Opaque Borrowers</i>	-0.034* (-1.77)	-0.035** (-1.97)	-0.013 (-0.67)
<i>Assets</i>	-0.008* (-1.88)	-0.007* (-1.69)	-0.008* (-1.75)
<i>Current Ratio</i>	-0.008** (-2.02)	-0.008** (-2.04)	-0.008** (-2.02)

TABLE 8 (continued)
Availability of Satellite Image Data and Alternative Source of Capital

Panel B: Availability of Satellite Image Data and Cost of Equity Capital—Opaque Borrowers (continued)

	AVG ICC		
	<i>No Analyst Coverage</i>	<i>No Earnings Forecast</i>	<i>Low Press Releases</i>
	(1)	(2)	(3)
<i>Leverage</i>	-0.009 (-0.44)	-0.008 (-0.41)	-0.008 (-0.42)
<i>MTB</i>	-0.001*** (-3.09)	-0.001*** (-3.05)	-0.001*** (-3.03)
<i>Sales Growth</i>	0.000 (0.02)	0.001 (0.08)	0.001 (0.09)
<i>Interest Coverage</i>	0.000 (0.07)	0.000 (0.10)	0.000 (0.03)
<i>ROA</i>	-0.049* (-1.75)	-0.048* (-1.71)	-0.052* (-1.81)
<i>Altman Z</i>	-0.001 (-0.54)	-0.000 (-0.44)	-0.001 (-0.50)
<i>AGE</i>	0.000 (1.13)	0.000 (1.01)	0.000 (1.07)
<i>Past Return</i>	0.002 (0.34)	0.001 (0.10)	0.001 (0.09)
<i>Amounts</i>	-0.001 (-0.64)	-0.001 (-0.67)	-0.001 (-0.51)
<i>Maturity</i>	-0.004 (-0.40)	-0.005 (-0.54)	-0.005 (-0.56)
<i>Secured</i>	0.004 (0.99)	0.004 (0.94)	0.005 (1.06)
<i>Guarantor</i>	0.005 (0.84)	0.005 (0.75)	0.005 (0.85)
<i>Model</i>	OLS	OLS	OLS
<i>Firm FE</i>	Yes	Yes	Yes
<i>Year FE</i>	No	No	No
<i>Year-Quarter FE</i>	Yes	Yes	Yes
<i>Credit Rating FE</i>	Yes	Yes	Yes
<i>Loan Type FE</i>	Yes	Yes	Yes
<i>Observations</i>	2,129	2,129	2,129
<i>Adj. (Pseudo) R-squared</i>	0.443	0.441	0.440
<i>Test: Treatment Firm x Post + Treatment Firm x Post x Opaque Borrowers = 0</i>	p-value: 0.8995	p-value: 0.3514	p-value: 0.3296

TABLE 8 (continued)
Availability of Satellite Image Data and Alternative Source of Capital

Panel C: Availability of the Satellite Image Data and Equity Issuance

	<i>Equity Issuance Amount</i>	<i>Equity Issuance Amount</i>	<i>Equity Issuance Indicator</i>	<i>Equity Issuance Indicator</i>	<i>Equity Issuance Indicator</i>
	(1)	(2)	(3)	(4)	(5)
<i>Treatment Firm x Post</i>	0.035 (0.11)	-0.041 (-0.13)	-0.747 (-0.75)	0.005 (0.14)	-0.002 (-0.05)
<i>Assets</i>	-0.304 (-1.41)	-0.394* (-1.96)	-0.212 (-0.60)	-0.042* (-1.66)	-0.052** (-2.19)
<i>Current Ratio</i>	-0.434*** (-2.71)	-0.476*** (-2.91)	-0.597** (-2.51)	-0.058*** (-2.83)	-0.063*** (-3.00)
<i>Leverage</i>	0.735 (1.20)	0.687 (1.13)	1.496 (1.57)	0.087 (1.10)	0.076 (0.98)
<i>MTB</i>	0.006 (0.81)	0.003 (0.43)	0.011 (0.63)	0.001 (0.66)	0.000 (0.27)
<i>Sales Growth</i>	0.190 (0.49)	0.048 (0.13)	0.305 (0.76)	0.031 (0.65)	0.016 (0.35)
<i>Interest Coverage</i>	0.000 (0.65)	0.000 (0.33)	0.001 (0.82)	0.000 (0.66)	0.000 (0.33)
<i>ROA</i>	0.738 (0.65)	0.758 (0.67)	0.829 (0.60)	0.062 (0.45)	0.056 (0.41)
<i>Altman Z</i>	-0.013 (-0.27)	0.004 (0.08)	-0.000 (-0.00)	0.000 (0.03)	0.002 (0.33)
<i>AGE</i>	-0.015 (-1.16)	-0.004 (-0.31)	-0.092* (-1.74)	-0.002 (-0.95)	-0.000 (-0.21)
<i>Past Return</i>	1.024*** (3.34)	0.927*** (2.93)	1.347*** (3.04)	0.124*** (3.29)	0.113*** (2.89)
<i>Amounts</i>	0.101 (0.88)	0.032 (0.31)	0.224 (1.03)	0.012 (0.86)	0.005 (0.42)
<i>Maturity</i>	-0.379 (-1.15)	-0.496 (-1.11)	-1.019 (-1.58)	-0.047 (-1.15)	-0.059 (-1.09)
<i>Secured</i>	0.185 (0.83)	0.186 (0.79)	0.205 (0.52)	0.019 (0.70)	0.018 (0.64)
<i>Guarantor</i>	0.399 (0.94)	0.280 (0.80)	0.499 (0.88)	0.056 (1.15)	0.044 (1.10)
<i>Model</i>	OLS	OLS	Logit	OLS	OLS
<i>Firm FE</i>	Yes	Yes	Yes	Yes	Yes
<i>Year FE</i>	Yes	No	Yes	Yes	No
<i>Year-Quarter FE</i>	No	Yes	No	No	Yes
<i>Credit Rating FE</i>	No	Yes	No	No	Yes
<i>Loan Type FE</i>	No	Yes	No	No	Yes
<i>Observations</i>	2,129	2,129	574	2,129	2,129
<i>Adj. (Pseudo) R-squared</i>	0.291	0.304	0.182	0.323	0.330

TABLE 8 (continued)
Availability of Satellite Image Data and Alternative Source of Capital

Panel D: Availability of the Satellite Image Data and Equity Issuance – Opaque Borrower

	<i>Equity Issuance Amount</i>			<i>Equity Issuance Indicator</i>		
	<i>No Analyst Coverage</i>	<i>No Earnings Forecast</i>	<i>Low Press Releases</i>	<i>No Analyst Coverage</i>	<i>No Earnings Forecast</i>	<i>Low Press Releases</i>
	(1)	(2)	(3)	(4)	(5)	(6)
<i>Treatment Firm x Post</i>	-0.039	-0.138	0.214	-0.008	-0.016	0.037
	(-0.12)	(-0.45)	(0.41)	(-0.21)	(-0.45)	(0.61)
<i>Post x Opaque Borrowers</i>	0.028	0.071	0.592	-0.011	0.003	0.070
	(0.08)	(0.18)	(1.61)	(-0.25)	(0.07)	(1.64)
<i>Treatment Firm x Post x Opaque Borrowers</i>	-0.001	1.197	-0.596	0.015	0.158	-0.087
	(-0.00)	(0.76)	(-0.91)	(0.19)	(0.80)	(-1.13)
<i>Assets</i>	-0.394*	-0.393*	-0.399**	-0.052**	-0.052**	-0.053**
	(-1.94)	(-1.94)	(-1.99)	(-2.15)	(-2.17)	(-2.22)
<i>Current Ratio</i>	-0.477***	-0.478***	-0.478***	-0.063***	-0.064***	-0.063***
	(-2.91)	(-2.93)	(-2.96)	(-3.00)	(-3.01)	(-3.04)
<i>Leverage</i>	0.684	0.703	0.680	0.077	0.077	0.075
	(1.12)	(1.17)	(1.13)	(0.99)	(1.01)	(0.98)
<i>MTB</i>	0.003	0.003	0.004	0.000	0.000	0.000
	(0.44)	(0.45)	(0.52)	(0.26)	(0.28)	(0.35)
<i>Sales Growth</i>	0.048	0.054	0.028	0.016	0.016	0.013
	(0.13)	(0.14)	(0.08)	(0.35)	(0.36)	(0.30)
<i>Interest Coverage</i>	0.000	0.000	0.000	0.000	0.000	0.000
	(0.33)	(0.33)	(0.23)	(0.31)	(0.33)	(0.21)
<i>ROA</i>	0.751	0.776	0.864	0.058	0.057	0.069
	(0.66)	(0.67)	(0.75)	(0.43)	(0.41)	(0.50)
<i>Altman Z</i>	0.004	0.004	0.008	0.002	0.002	0.003
	(0.08)	(0.07)	(0.15)	(0.34)	(0.31)	(0.39)
<i>AGE</i>	-0.004	-0.004	-0.004	-0.000	-0.000	-0.000
	(-0.31)	(-0.31)	(-0.36)	(-0.22)	(-0.21)	(-0.26)

TABLE 8 (continued)
Availability of Satellite Image Data and Alternative Source of Capital

Panel D: Availability of the Satellite Image Data and Equity Issuance – Opaque Borrower (continued)

	<i>Equity Issuance Amount</i>			<i>Equity Issuance Indicator</i>		
	<i>No Analyst Coverage</i>	<i>No Earnings Forecast</i>	<i>Low Press Releases</i>	<i>No Analyst Coverage</i>	<i>No Earnings Forecast</i>	<i>Low Press Releases</i>
	(1)	(2)	(3)	(4)	(5)	(6)
<i>Past Return</i>	0.925*** (2.91)	0.915*** (2.88)	0.921*** (2.89)	0.113*** (2.88)	0.111*** (2.84)	0.112*** (2.86)
<i>Amounts</i>	0.032 (0.31)	0.034 (0.33)	0.028 (0.28)	0.005 (0.41)	0.006 (0.45)	0.005 (0.39)
<i>Maturity</i>	-0.498 (-1.12)	-0.503 (-1.13)	-0.506 (-1.14)	-0.059 (-1.08)	-0.060 (-1.11)	-0.061 (-1.12)
<i>Secured</i>	0.186 (0.79)	0.184 (0.79)	0.170 (0.73)	0.018 (0.64)	0.018 (0.63)	0.016 (0.59)
<i>Guarantor</i>	0.280 (0.80)	0.275 (0.78)	0.289 (0.83)	0.044 (1.11)	0.044 (1.09)	0.045 (1.14)
<i>Model</i>	OLS	OLS	OLS	OLS	OLS	OLS
<i>Firm FE</i>	Yes	Yes	Yes	Yes	Yes	Yes
<i>Year FE</i>	No	No	No	No	No	No
<i>Year-Quarter FE</i>	Yes	Yes	Yes	Yes	Yes	Yes
<i>Credit Rating FE</i>	Yes	Yes	Yes	Yes	Yes	Yes
<i>Loan Type FE</i>	Yes	Yes	Yes	Yes	Yes	Yes
<i>Observations</i>	2,129	2,129	2,129	2,129	2,129	2,129
<i>Adj. (Pseudo) R-squared</i>	0.303	0.304	0.305	0.329	0.330	0.331

TABLE 9*Institutional Lenders' Information Demand and Pricing Dynamics in the Institutional Loan Market*

This table examines whether institutional lenders' information demand, particularly for opaque borrowers, affects interest spreads in the pre-coverage period. Across columns, I use the following variables as proxies for borrower opacity: *No Analyst Coverage* in Columns 1 and 2, *No Earnings Forecast* in Columns 3 and 4, and *Low Press Releases* in Columns 5 and 6. t-statistics in parentheses are based on standard errors clustered at the firm level. ***, **, and * indicate significance at the 0.01, 0.05, and 0.10 level, respectively. All variables are defined in Appendix A.

	<i>Interest Spread</i>					
	<i>No Analyst Coverage</i>		<i>No Earnings Forecast</i>		<i>Low Press Releases</i>	
	(1)	(2)	(3)	(4)	(5)	(6)
<i>Inst. Lender</i>	0.010 (0.49)	0.002 (0.14)	0.063** (2.15)	0.059** (2.32)	0.014 (0.45)	0.005 (0.23)
<i>Borrower Opacity</i>	-0.017 (-0.52)	-0.013 (-0.43)	0.036 (1.57)	0.045** (2.07)	0.023 (0.93)	0.001 (0.03)
<i>Inst. Lender x Borrower Opacity</i>	0.108** (2.44)	0.067* (1.75)	-0.051 (-1.04)	-0.072 (-1.60)	0.068 (1.42)	0.046 (1.13)
<i>Assets</i>	-0.006 (-0.74)	0.000 (0.03)	-0.004 (-0.53)	0.001 (0.13)	-0.009 (-0.90)	0.001 (0.07)
<i>Current Ratio</i>	0.011 (1.16)	0.005 (0.53)	0.012 (1.25)	0.011 (1.10)	0.010 (0.91)	0.005 (0.52)
<i>Leverage</i>	0.219*** (5.12)	0.135*** (3.78)	0.223*** (5.26)	0.172*** (4.42)	0.266*** (5.73)	0.136*** (3.82)
<i>MTB</i>	-0.001 (-1.09)	-0.001 (-1.57)	-0.001 (-1.12)	-0.001*** (-2.68)	-0.001** (-2.20)	-0.001* (-1.69)
<i>Sales Growth</i>	-0.012 (-0.51)	0.013 (0.57)	-0.012 (-0.48)	-0.005 (-0.22)	-0.017 (-0.70)	0.016 (0.70)
<i>Interest Coverage</i>	-0.000 (-0.18)	-0.000 (-0.66)	-0.000 (-0.20)	-0.000 (-0.85)	-0.000 (-0.06)	-0.000 (-0.62)
<i>ROA</i>	-0.066 (-0.74)	0.003 (0.04)	-0.043 (-0.49)	0.047 (0.58)	-0.009 (-0.09)	0.015 (0.20)
<i>Altman Z</i>	-0.014*** (-3.70)	-0.016*** (-4.44)	-0.015*** (-3.73)	-0.009** (-1.97)	-0.011** (-2.29)	-0.016*** (-4.54)
<i>AGE</i>	-0.001* (-1.84)	-0.000 (-0.85)	-0.001* (-1.80)	-0.000 (-0.00)	-0.001 (-1.64)	-0.001 (-0.89)
<i>Past Return</i>	0.011 (0.33)	0.008 (0.28)	0.013 (0.39)	0.005 (0.18)	-0.009 (-0.28)	0.010 (0.37)
<i>AGE</i>	-0.001* (-1.84)	-0.000 (-0.85)	-0.001* (-1.80)	-0.000 (-0.00)	-0.001 (-1.64)	-0.001 (-0.89)
<i>Past Return</i>	0.011 (0.33)	0.008 (0.28)	0.013 (0.39)	0.005 (0.18)	-0.009 (-0.28)	0.010 (0.37)

TABLE 9 (continued)*Institutional Lenders' Information Demand and Pricing Dynamics in the Institutional Loan Market*

	<i>Interest Spread</i>					
	<i>No Analyst Coverage</i>		<i>No Earnings Forecast</i>		<i>Low Press Releases</i>	
	(1)	(2)	(3)	(4)	(5)	(6)
<i>Amounts</i>	-0.076*** (-6.30)	-0.083*** (-7.60)	-0.081*** (-6.68)	-0.084*** (-6.38)	-0.070*** (-4.93)	-0.083*** (-7.65)
<i>Maturity</i>	-0.153*** (-3.82)	-0.134*** (-3.17)	-0.146*** (-3.62)	-0.095** (-2.11)	-0.168*** (-3.78)	-0.135*** (-3.21)
<i>Secured</i>	0.216*** (11.03)	0.106*** (5.70)	0.220*** (11.24)	0.113*** (5.58)	0.202*** (9.36)	0.108*** (5.76)
<i>Guarantor</i>	-0.109*** (-4.18)	-0.048** (-2.05)	-0.105*** (-4.13)	-0.051** (-2.20)	-0.100*** (-3.67)	-0.047** (-2.04)
<i>Model</i>	OLS	OLS	OLS	OLS	OLS	OLS
<i>Firm FE</i>	No	No	No	No	No	No
<i>Year FE</i>	No	No	No	No	No	No
<i>Year-Quarter FE</i>	Yes	Yes	Yes	Yes	Yes	Yes
<i>Credit Rating FE</i>	No	Yes	No	Yes	No	Yes
<i>Loan Type FE</i>	No	Yes	No	Yes	No	Yes
<i>Observations</i>	1,422	1,422	1,422	1,422	1,422	1,422
<i>Adj. (Pseudo) R-squared</i>	0.355	0.522	0.352	0.556	0.400	0.521